# PRICING AND HEDGING CONTINGENT CLAIMS IN A MULTI-ASSET BINOMIAL MARKET 

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#### Abstract

We consider an incomplete multi-asset binomial market model. We prove that for a wide class of contingent claims the extremal multi-step martingale measure is a power of the corresponding single-step extremal martingale measure. This allows for closed form formulas for the bounds of a no-arbitrage contingent claim price interval. We construct a feasible algorithm for computing those boundaries as well as for the corresponding hedging strategies. Our results apply, for example, to European basket call and put options and Asian arithmetic average options.


## 1. Introduction

We consider a discrete time $n$-step market model with $m$ assets each of which is following a binomial model. We consider both path-dependent and path-independent multi-asset contingent claims. Since the model is incomplete for $m \geq 2$, there is no unique no-arbitrage price of a contingent claim, but rather an interval of no-arbitrage prices. We compute the bounds of that interval and provide an algorithm for such computations. We also construct corresponding extremal hedging strategies.

The computations amount to extremizing the expected pay-off of a contingent claim over the set of appropriate martingale measures. In general, such a set is not easy to understand, although it is the convex hull of finitely many points in a vector space. Our main results state that for certain contingent claims the extrema of no-arbitrage prices are attained at martingale measures of the form $\mathbf{P}^{\otimes n}=\mathbf{P} \otimes \mathbf{P} \otimes \cdots \otimes \mathbf{P}$, where $\mathbf{P}$ is a 1-step martingale measure. This remarkable fact allows writing down explicit formulas and constructing effective algorithms.

Let us state the main results precisely. Let $S_{i}(j)$, where $i=1, \ldots, m$ and $j=0, \ldots, n$, denote the price of the $i$-th asset at time $j$. Let $\left(S_{i}(j)\right)$ denote a random element of $\mathbb{R}^{m n}$ and let $X=f\left(S_{i}(j)\right)$ be a contingent claim, where $f: \mathbb{R}^{m n} \rightarrow \mathbb{R}$ is a function. We prove (Proposition 4.1) that there exist measures $\mathbf{P}_{\text {max }}$ and $\mathbf{P}_{\text {min }}$ such that the upper and lower bounds of the no-arbitrage price interval of $X$ at each time are given by the (discounted) conditional expectation of $X$ with respect to $\mathbf{P}_{\text {max }}$ and $\mathbf{P}_{\text {min }}$, respectively. In general, finding these extremal measures requires running a linear program a number of times with grows like $2^{m n}$. This is because they are built from extremal single-step martingale measures and at each time such a measure has to be found by solving a linear programming problem. In what follows we present the results that require running the linear program once. This is because the extremal measure is a product of the same single-step martingale measure so it is enough to solve a linear programming problem just once. The following result is a special case of a slightly more general Theorem 5.4 proven in Section 5B.

Theorem 1.1. Let $X$ be a contingent claim which as a random variable is fibrewise supermodular (see Definition A.10). Then there exists a 1 -step martingale measure $\mathbf{P}$ such that $\mathbf{P}_{\max }=\mathbf{P}^{\otimes n}$. If, moreover, the stock price ratios satisfy

$$
\sum_{i} \frac{R-D_{i}}{U_{i}-D_{i}}<1
$$

then also the minimal martingale measure is a product on $n$ copies of a 1-step martingale measure.

Contingent claims satisfying the hypothesis of the above theorem include European basket call and put options (Example 7.3 and 7.4), Asian (path-dependent) basket options based on arithmetic average (Example 7.5). For European basket call options we present explicit formulae for the bounds of no-arbitrage values in Example 7.3. Although the theorem does not apply to Asian contingent claims based on geometric average, we can still provide estimates of no-arbitrage prices (Example 7.7).

Since Theorem 1.1 yields relatively fast computation of the extremal no-arbitrage values of contingent claims, it also allows constructing effective algorithms for computing extremal hedging strategies. We discuss them in Sections 3F, 4D and 6C.

Remark 1.2. The proof of Theorem 1.1 is an elaboration of a well known result of combinatorial optimization which shows that supermodular set functions when restricted to convex polytopes are maximized on a special vertex [1, 5, 6, 7]. We present detailed arguments in the Appendix, which is somewhat more technical than the rest of the paper.

Remark 1.3. The case of two assets has been solved in the paper by Nagaev and Steblovskaya [3], where they also consider the case of continuous distribution of prices. A geometric approach to the problem was taken by Motoczyński and Stettner [2].

Structure of the paper. Section 2 is devoted to introducing all necessary definitions. We put an emphasis on working with concrete sample spaces for several reasons. Namely, one of our aims is to create a workable algorithm and a suitable computer program. Moreover, our proofs then deal with concrete polytopes and concrete optimization problems. In Section 3 we discuss 1-step model that constitutes building blocks for the multi-step model discussed in Section 4, where we also discuss hedging strategies. Algorithms and the main theorem are discussed in Section 5. In Section 6, we provide explicit pricing formulas and an algorithm for an extremal hedging strategy. Finally, in Section 7, we discuss concrete examples of contingent claims. All technical results needed for the proofs are presented in the Appendix.

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## 2. Preliminaries

2A. The market model and main definitions. We consider a financial market which consists of $m$ risky assets (stocks) and a riskless asset (bond) that can be traded at discrete
time moments $t=0,1, \ldots, n$. In this market we consider a European contingent claim (option).

Let us define the main ingredients of the model in more detail.
Time. The time is modeled by a finite set $\mathbb{T}=\{0,1,2, \ldots, n\}$.
The bond process. It is a deterministic process defined by $B(t)=R^{t}$, where $R=1+r$ and $r$ is a constant periodic interest rate.

The stock price process. For each $i=1, \ldots, m$, the stock price process $S_{i}=\left(S_{i}(t)\right)_{t=0,1, \ldots, n}$ is described by the $n$-step binomial dynamics. The stock price ratios $\psi_{i}(t)=S_{i}(t) / S_{i}(t-1)$ at each time moment $t=1, \ldots, n$ can take two possible values: $\psi_{i}(t) \in\left\{D_{i}, U_{i}\right\}$. We will assume that for each binomial model the no-arbitrage condition $0<D_{i}<R<U_{i}$ holds.

In order to simplify terminology we say that the price of the $i$-th stock went $\boldsymbol{u p}$ at time $t$ if $\psi_{i}(t)=U_{i}$ and it went down if $\psi_{i}(t)=D_{i}$. We will denote by $S(t)$ a stock price vector at time $t$ :

$$
S(t)=\left(S_{1}(t), S_{2}(t), \ldots, S_{m}(t)\right),
$$

where $t=0,1, \ldots, n$.
The contingent claim. In the above market, we consider a European contingent claim $X$ with the real-valued pay-off function $f$, which may depend only on the terminal value $S(n)$ of the stock price vector or on the entire stock price path $S(t), t=1, \ldots, n$.

If $m \geq 2$ then the above market model is incomplete [4, Section 1.5].
Sample space for the $n$-step model and basic random variables. The sample space $\Omega_{n}$ for the $n$-step model is modeled by the set of $(m \times n)$-matrices $\omega$ with entries $\omega_{i j} \in\{0,1\}$. Each element $\omega$ represents the state of the world at time $n$. The value $\omega_{i j}=0$ signifies that the price of the $i$-th stock went down at time $j$. The value $\omega_{i j}=1$ signifies that the price of the $i$-th stock went up at time $j$.
With this notation, the stock price ratio $\psi_{i}(t)=S_{i}(t) / S_{i}(t-1)$ for each $t=1, \ldots, n$ can be defined as a random variable on $\Omega_{n}$ described as follows:

$$
\psi_{i}(t)(\omega)= \begin{cases}D_{i} & \text { if } \omega_{i t}=0 \\ U_{i} & \text { if } \omega_{i t}=1\end{cases}
$$

where $i=1, \ldots, m$. Consequently, the $i$-th stock price at time $t \in \mathbb{T}$ can be presented as

$$
S_{i}(t)=S_{i}(0) \psi_{i}(1) \cdots \psi_{i}(t)
$$

where $i=1, \ldots, m$ and $S_{i}(0)$ is the known initial stock price.
A direct computation shows that

$$
S_{i}(t)(\omega)=S_{i}(0) U_{i}^{\sum_{j=1}^{t} \omega_{i j}} D_{i}^{t-\sum_{j=1}^{t} \omega_{i j}}
$$

So for each $\omega \in \Omega_{n}$ the $i$-th row of $\omega$ describes the $n$-step dynamics of the $i$-th stock price, while the $j$-th column describes the single-step dynamics of the stock price vector from time $j-1$ to time $j$.

Ordering elements in a sample space. Whenever convenient we shall consider an elementary event $\omega \in \Omega_{n}$ either as a matrix

$$
\omega=\left(\begin{array}{cccc}
\omega_{11} & \omega_{12} & \ldots & \omega_{1 n} \\
\omega_{21} & \omega_{22} & \ldots & \omega_{2 n} \\
\vdots & \vdots & \ldots & \vdots \\
\omega_{m 1} & \omega_{m 2} & \ldots & \omega_{m n}
\end{array}\right)
$$

or as an $n$-tuple of its column vectors

$$
\omega=\left(\omega^{1} \omega^{2} \ldots \omega^{n}\right)
$$

where

$$
\omega^{j}=\left(\begin{array}{c}
\omega_{1 j} \\
\omega_{2 j} \\
\vdots \\
\omega_{m j}
\end{array}\right)
$$

It is straightforward to see that $\Omega_{n}$ consists of $N=2^{m n}$ elements.
It will be convenient to order the elements of $\Omega_{n}$ with respect to the reverse lexicographic order. In the case of $n=1$, the set $\Omega_{n}=\Omega_{1}$ contains $N=2^{m}$ elements. Each element $\omega \in \Omega_{1}$ is a ( $m \times 1$ )-matrix (or, equivalently, a column vector of length $m$ ), and the elements of $\Omega_{1}$ are ordered as follows:

$$
\omega_{1}=\left(\begin{array}{c}
1  \tag{2.1}\\
1 \\
\vdots \\
1 \\
1
\end{array}\right), \omega_{2}=\left(\begin{array}{c}
1 \\
1 \\
\vdots \\
1 \\
0
\end{array}\right), \ldots, \omega_{N-1}=\left(\begin{array}{c}
0 \\
0 \\
\vdots \\
0 \\
1
\end{array}\right), \omega_{N}=\left(\begin{array}{c}
0 \\
0 \\
\vdots \\
0 \\
0
\end{array}\right) .
$$

Each $\omega_{i}$, when read from top to bottom, is a binary representation of the number $2^{m}-i$.
For the case $n=2$, the set $\Omega_{n}=\Omega_{2}$ contains $N=2^{m \times 2}$ elements. Each element $\omega \in \Omega_{2}$ is a ( $m \times 2$ )-matrix, or, equivalently, a pair of column vectors of length $m$ :

$$
\omega=\left(\omega^{1} \omega^{2}\right)
$$

The elements of $\Omega_{2}$ are ordered as follows. In the first $2^{m}$ elements of $\Omega_{2}$, the first column $\omega^{1}$ is fixed at $\omega_{1}$, where $\omega_{1}$ is defined in (2.1), and the second column $\omega^{2}$ runs through the set $\Omega_{1}$ according to the order in (2.1). In the next $2^{m}$ elements of $\Omega_{2}$, the first column $\omega^{1}$ is fixed at $\omega_{2}$ and the second column $\omega^{2}$ runs through the set $\Omega_{1}$ according to the order in (2.1). Continuing this way, one orders all elements in the set $\Omega_{2}$.

Elements of $\Omega_{n}$ for any $n$ can be ordered in a similar manner.

The information structure. The state of the world at time $k \in \mathbb{T}$ is described by the subset of matrices from $\Omega_{n}$ with the first $k$ columns fixed. Each subset, denoted $\mathcal{P}\left(\omega^{1}, \ldots, \omega^{k}\right)$, has
the following form:

$$
\begin{align*}
& \mathcal{P}\left(\omega^{1}, \ldots, \omega^{k}\right)=\left\{\omega \in \Omega_{n} \left\lvert\, \omega=\left(\begin{array}{llllll}
\omega^{1} & \ldots & \omega^{k} & * & \ldots & *
\end{array}\right)\right.\right\} \\
& =\left\{\omega \in \Omega_{n} \left\lvert\, \omega=\left(\begin{array}{cccccc}
\omega_{11} & \ldots & \omega_{1 k} & * & \ldots & * \\
\omega_{21} & \ldots & \omega_{2 k} & * & \ldots & * \\
\vdots & \vdots & \vdots & \vdots & & \\
\omega_{m 1} & \ldots & \omega_{m k} & * & \ldots & *
\end{array}\right)\right.\right\} \tag{2.2}
\end{align*}
$$

In what follows, we will say that the $(m \times k)$-matrix $\left(\omega^{1} \ldots \omega^{k}\right)$ which represents the common part of all matrices $\omega \in \Omega_{n}$ included in set $\mathcal{P}\left(\omega^{1}, \ldots, \omega^{k}\right)$ is associated with the set $\mathcal{P}\left(\omega^{1}, \ldots, \omega^{k}\right)$.

There are $2^{m k}$ disjoint subsets of the form (2.2) with different associated matrices. These subsets form a partition $\mathcal{P}_{k}$ of $\Omega_{n}$. The initial partition is trivial $\mathcal{P}_{0}=\left\{\Omega_{n}\right\}$, the last one $\mathcal{P}_{n}=\left\{\left\{\omega_{1}\right\},\left\{\omega_{2}\right\}, \ldots,\left\{\omega_{N}\right\}\right\}$, where $N=2^{m n}$. Clearly, the partition $\mathcal{P}_{k}$ is finer than the partition $\mathcal{P}_{k-1}$ and hence they form a sequence of finer and finer partitions.

Each partition $\mathcal{P}_{k}$ can be put into one-to-one correspondence with a subalgebra $\mathcal{F}_{k}$ of the algebra $2^{\Omega_{n}}$ of all subsets of $\Omega_{n}$. The subalgebras $\mathcal{F}_{k}$ form a filtration $\mathcal{F}$, an increasing sequence of subalgebras $\left\{\mathcal{F}_{k}\right\}, k=0, \ldots, n-1$, where $\mathcal{F}_{k} \subseteq \mathcal{F}_{k+1}$. Here $\mathcal{F}_{0}=\left\{\emptyset, \Omega_{n}\right\}$, $\mathcal{F}_{n}=2^{\Omega_{n}}$ consists of all subsets of $\Omega_{n}$.

In what follows, we assume that the filtration $\mathcal{F}$ is generated by the stock price vector process $(S(t))_{t=0,1, \ldots, n}$.

The supporting tree. The above information structure can be described also with the help of a finite directed rooted $n$-step binary tree which we will denote by $\mathbf{T}$. We will call $\mathbf{T}$ the supporting tree for the $n$-step market model under consideration. The supporting tree $\mathbf{T}$ consists of vertices and directed edges that connect the vertices.

In what follows, we will use the following terminology and notation. We will denote by $v_{0} \in \mathbf{T}$ the root of the tree $\mathbf{T}$. In other words, $v_{0}$ is the vertex that corresponds to time $t=0$.

For each vertex $v \in \mathbf{T}$ there is a unique path from the root $v_{0}$ to $v$. If such a path consists of $k$ edges we say that the vertex $v$ corresponds to the time step $k$. The set of vertices that correspond to time $k$ is denoted by $\mathbf{T}_{k}$. In particular, $\mathbf{T}_{n}$ are the terminal vertices. We will call terminal vertices leaves and non-terminal vertices nodes. In other words, $v \in \mathbf{T}_{k}$ is called a leaf if $k=n$. If $k<n, v$ is called a node. The root $v_{0}$ of the tree is a special node that corresponds to $k=0$.

We say that a vertex $v \in \mathbf{T}_{\ell}, \ell \leq n$, is a successor of a node $w \in \mathbf{T}_{k}, k<\ell$, if there is a path from $w$ to $v$.

The leaves of the tree $\mathbf{T}$ are in one-to-one correspondence with the elements of the sample space $\Omega_{n}$, or, in other words, with the sets of the partition $\mathcal{P}_{n}=\left\{\left\{\omega_{1}\right\},\left\{\omega_{2}\right\}, \ldots,\left\{\omega_{N}\right\}\right\}$. The set of leaves corresponds to the time step $n$ and represents all possible states of the world at time $n$. Additionally, each leaf describes the stock price vector dynamics from time 0 to time $n$.

The nodes of $\mathbf{T}$ that correspond to time $k<n$ are in one-to-one correspondence with the sets $\mathcal{P}\left(\omega^{1}, \ldots, \omega^{k}\right)$ of the partition $\mathcal{P}_{k}$ that are defined in (2.2).

In the rest of the paper we will frequently use this correspondence. So whenever we are talking about a node $v \in \mathbf{T}_{k}$ we implicitly mean that it is the corresponding set $\mathcal{P}\left(\omega^{1}, \ldots, \omega^{k}\right)$ of the partition $\mathcal{P}_{k}$. The $(m \times k)$-matrix $\left(\omega^{1} \ldots \omega^{k}\right)$ associated with the set $\mathcal{P}\left(\omega^{1}, \ldots, \omega^{k}\right)$ will be called also the matrix associated with the node $v \in \mathbf{T}_{k}$.

Similarly to the case of leaves, each node $v \in \mathbf{T}_{k}, k<n$, describes the stock price vector dynamics from time 0 to time $k$.

There is an edge from a node $v$ to a vertex $w$ (where $w$ could be a node or a leaf) if and only if a matrix associated with $w$ has been obtained from the matrix associated $v$ by appending a column. Thus the root is associated with an empty matrix and the tree is regular in the sense that for each vertex except the leaves and the root there is exactly one incoming edge and $2^{m}$ outgoing edges.

Probability measures in an n-step model. A probability measure $\mathbf{P}$ on $\left(\Omega_{n}, 2^{\Omega_{n}}, \mathcal{F}\right)$ (or, equivalently, a probability measure in an $n$-step market model) is defined by its probability function $p: \Omega_{n} \rightarrow[0,1]$ :

$$
p\left(\omega_{i}\right)=\mathbf{P}\left(\left\{\omega_{i}\right\}\right)=p_{i},
$$

where $\omega_{i} \in \Omega_{n}$ is an elementary event, $i=1, \ldots, N=2^{m n}$, and $\sum_{i=1}^{N} p_{i}=1$. For simplicity, in what follows we will use the notation $\mathbf{P}\left(\omega_{i}\right)$ instead of $\mathbf{P}\left(\left\{\omega_{i}\right\}\right)$ and identify a probability measure $\mathbf{P}$ with a vector $\left(p_{1}, p_{2}, \ldots, p_{N}\right) \in \mathbb{R}^{N}$ with non-negative coordinates which sum up to 1 . Thus the set of all probability measures in the $n$-step market model with $m$ assets is equivalent to the unit simplex:

$$
\begin{equation*}
\Delta\left(\Omega_{n}\right)=\left\{\left(p_{1}, \ldots, p_{N}\right) \in \mathbb{R}^{N} \mid \sum_{i=1}^{N} p_{i}=1, p_{i} \geq 0, N=2^{m n}\right\} \tag{2.3}
\end{equation*}
$$

In what follows we will use the notation $\Delta\left(\Omega_{n}\right)$ for the set of probability measures on $\left(\Omega_{n}, 2^{\Omega_{n}}, \mathcal{F}\right)$.

Martingale and risk-neutral measures. A probability measure $\mathbf{P} \in \Delta\left(\Omega_{n}\right)$ is called a martingale measure in an $n$-step market model if it satisfies the following conditions:

$$
\begin{equation*}
\mathbf{E}_{\mathbf{P}}\left(S_{i}(k+\ell) \mid \mathcal{F}_{k}\right)=R^{\ell} S_{i}(k), \tag{2.4}
\end{equation*}
$$

or equivalently,

$$
\begin{equation*}
\mathbf{E}_{\mathbf{P}}\left(\psi_{i}(k+1) \ldots \psi_{i}(k+\ell)\right)=R^{\ell} \tag{2.5}
\end{equation*}
$$

where $i=1,2, \ldots, m$ and $0 \leq k+\ell \leq n$ with $k, \ell \geq 0$. In other words, $\mathbf{P}$ is a martingale measure if and only if the discounted price process for each stock is a martingale with respect to $\mathbf{P}$. The set of martingale measures in an $n$-step market model will be denoted by $\mathrm{M}_{n}$.

A martingale measure $\mathbf{P}=\left(p_{1}, p_{2}, \ldots, p_{N}\right)$ is called risk-neutral if $p_{i}>0$ for each $i=$ $1, \ldots, N$. The set of risk-neutral measures in an $n$-step market model will be denoted by $\mathrm{N}_{n}$. So we have: $\mathrm{N}_{n} \subset \mathrm{M}_{n} \subset \Delta\left(\Omega_{n}\right)$, where $\mathrm{M}_{n}$ is the closure of $\mathrm{N}_{n}$ in $\mathbb{R}^{N}$.

2B. Supporting known results. In this section we present a number of known results which we need later.

Multi-step vs single-step measures. Let $\mathbf{P} \in \Delta\left(\Omega_{n}\right)$ be a probability measure. Let us assume in addition $\mathbf{P}$ is non-degenerate in the following sense: $\mathbf{P}\left(\omega_{i}\right)=p_{i}>0$ for each $i=1, \ldots, N$. For each node $v \in \mathbf{T}_{k}, k<n$, the $n$-step non-degenerate probability measure $\mathbf{P}$ defines a non-degenerate probability measure $\mathbf{P}_{v} \in \Delta\left(\Omega_{1}\right)$ in the corresponding underlying single-step model, as the following conditional probability given the node $v$ :

$$
\begin{equation*}
\mathbf{P}_{v}\left(\omega^{k+1}\right)=\mathbf{P}\left(\mathcal{P}\left(\omega^{1} \cdots \omega^{k} \omega^{k+1}\right) \mid \mathcal{P}\left(\omega^{1} \cdots \omega^{k}\right)\right) \tag{2.6}
\end{equation*}
$$

for $k=0, \ldots, n-1$. Here $\mathcal{P}\left(\omega^{1} \ldots \omega^{k}\right)$ is the set of the partition $\mathcal{P}_{k}$ corresponding to the node $v \in \mathbf{T}_{k}$, and $\omega^{k+1}$ is an element of the sample space $\Omega_{1}$ in the corresponding underlying single-step model with the root at node $v$. Since $\mathbf{P}$ is non-degenerate, the event $\mathcal{P}\left(\omega^{1} \ldots \omega^{k}\right)$ has positive measure and the above conditional probability is well defined.

Conversely, assigning a single-step, possibly degenerate, probability $\mathbf{P}_{v} \in \Delta\left(\Omega_{1}\right)$ at each node $v \in \mathbf{T}_{k}, k<n$, defines an $n$-step probability measure $\mathbf{P} \in \Delta\left(\Omega_{n}\right)$ as follows: for each $\omega=\left(\omega^{1} \ldots \omega^{n}\right) \in \Omega_{n}$,

$$
\begin{equation*}
\mathbf{P}(\omega)=\mathbf{P}_{v_{0}}\left(\omega^{1}\right) \mathbf{P}_{v_{1}}\left(\omega^{2}\right) \cdots \mathbf{P}_{v_{n-1}}\left(\omega^{n}\right) \tag{2.7}
\end{equation*}
$$

where the node $v_{k}$ corresponds to the set $\mathcal{P}\left(\omega^{1} \cdots \omega^{k}\right)$.
Multi-step and single-step martingale and risk-neutral measures. The above construction preserves the martingale and risk-neutral properties of measures as stated in the following proposition the proof of which can be found in [4, Section 3.4].

Proposition 2.1. Let $\mathbf{P} \in \Delta\left(\Omega_{n}\right)$ and let $\mathbf{T}$ be the supporting tree. Then $\mathbf{P}$ is a martingale measure (resp. risk-neutral measure) in an n-step market model if and only if each $\mathbf{P}_{v_{k}} \in$ $\Delta\left(\Omega_{1}\right)$ in (2.7) is a martingale measure (resp. risk-neutral measure) in the corresponding underlying single-step model. In other words:

$$
\begin{aligned}
& \mathbf{P} \in \mathrm{M}_{n} \Leftrightarrow \forall k \mathbf{P}_{v_{k}} \in \mathrm{M}_{1} \\
& \mathbf{P} \in \mathrm{~N}_{n} \Leftrightarrow \forall k \mathbf{P}_{v_{k}} \in \mathrm{~N}_{1} .
\end{aligned}
$$

No-arbitrage pricing of contingent claims. Let $X$ be a European type contingent claim in an $n$-step market model. Since our market model is incomplete for $m \geq 2$, the no-arbitrage price of $X$ at time $k=0,1, \ldots, n-1$ is not unique. Each no-arbitrage price of $X$ at time $k$ is obtained as a discounted conditional expectation with respect to an $n$-step risk-neutral measure $\mathbf{P} \in \mathrm{N}_{n}$ as follows:

$$
\begin{equation*}
\mathrm{C}_{\mathbf{P}}(X, k)=R^{-(n-k)} \mathbf{E}_{\mathbf{P}}\left(X \mid \mathcal{F}_{k}\right) . \tag{2.8}
\end{equation*}
$$

Notice that for $k>0, \mathrm{C}_{\mathbf{P}}(X, k)$ is a random variable measurable with respect to the algebra $\mathcal{F}_{k}$ and hence it is determined by its values on the sets of the partition $\mathcal{P}_{k}$, that is, on the nodes $v \in \mathbf{T}_{k}$ of the supporting tree at time $k$. By varying the risk-neutral measures we obtain that for each $v \in \mathbf{T}_{k}$ the set of no-arbitrage prices of $X$ is an open interval:

$$
\begin{equation*}
\left(\mathrm{C}_{\min }(v), \mathrm{C}_{\max }(v)\right)=\left\{\mathrm{C}_{\mathbf{P}}(X, k)(v) \in \mathbb{R} \mid \mathbf{P} \in \mathrm{N}_{n}\right\} . \tag{2.9}
\end{equation*}
$$

## 3. The single-step case

In this section we consider a single-step model, that is, we assume $n=1$ throughout.

3A. Specification of the model. Recall from Section 2 that the sample space $\Omega_{1}$ for the single-step model consists of $(m \times 1)$-matrices with binary coefficients. Thus $\Omega_{1}$ has $N=2^{m}$ elements and they are ordered with respect to the reverse lexicographical order. For example, if $m=3$ then we obtain the following sequence of elements of $\Omega_{1}$ :

$$
\left(\begin{array}{l}
1 \\
1 \\
1
\end{array}\right)\left(\begin{array}{l}
1 \\
1 \\
0
\end{array}\right)\left(\begin{array}{l}
1 \\
0 \\
1
\end{array}\right)\left(\begin{array}{l}
1 \\
0 \\
0
\end{array}\right)\left(\begin{array}{l}
0 \\
1 \\
1
\end{array}\right)\left(\begin{array}{l}
0 \\
1 \\
0
\end{array}\right)\left(\begin{array}{l}
0 \\
0 \\
1
\end{array}\right)\left(\begin{array}{l}
0 \\
0 \\
0
\end{array}\right)
$$

There are only two instances of time $\mathbb{T}=\{0,1\}$ and the initial prices $S_{i}(0)$ of the risky assets are known; here $i=1,2, \ldots, m$. The terminal prices are random variables given by $S_{i}(1)=S_{i}(0) \psi_{i}(1)$, where $\psi_{i}(1) \in\left\{D_{i}, U_{i}\right\}$ are price ratios.

3B. Martingale measures. The system of equations (2.5) which defines the set $\mathrm{M}_{n}$ of martingale measures on $\Omega_{n}$ takes the following form for $n=1$ :

$$
\mathbf{E}_{\mathbf{P}}\left(\psi_{i}(1)\right)=R,
$$

where $i=1, \ldots, m$. It follows that a measure $\mathbf{P}=\left(p_{1}, p_{2}, \ldots, p_{N}\right)$ on $\Omega_{1}$ is a martingale measure $\left(\mathbf{P} \in \mathrm{M}_{1}\right)$ if it satisfies the following system of equations and inequalities:

$$
\begin{array}{rlll}
\psi_{1}\left(\omega_{1}\right) p_{1}+\psi_{1}\left(\omega_{2}\right) p_{2}+\cdots+\psi_{1}\left(\omega_{N}\right) p_{N} & = & R  \tag{3.1}\\
\cdots & \cdots & \\
\psi_{m}\left(\omega_{1}\right) p_{1}+\psi_{m}\left(\omega_{2}\right) p_{2}+\cdots+\psi_{m}\left(\omega_{N}\right) p_{N} & = & R \\
p_{1}+p_{2}+\cdots+p_{N} & =1 \\
p_{j} & \geq 0 &
\end{array}
$$

Here we used the simplified notation $\psi_{i}\left(\omega_{j}\right)$ to denote $\psi_{i}(1)\left(\omega_{j}\right)$.
Let $\boldsymbol{\Psi}$ be an $(m \times N)$-matrix that corresponds to the first $m$ equations in the above system. That is,

$$
\begin{equation*}
\boldsymbol{\Psi}_{i j}=\psi_{i}\left(\omega_{j}\right) \tag{3.2}
\end{equation*}
$$

for $i=1, \ldots, m$ and $j=1, \ldots, N=2^{m}$. Taking into account the ordering of $\Omega_{1}$ we obtain, as an example for $m=3$, the following expression for $\boldsymbol{\Psi}$ :

$$
\Psi=\left(\begin{array}{cccccccc}
U_{1} & U_{1} & U_{1} & U_{1} & D_{1} & D_{1} & D_{1} & D_{1} \\
U_{2} & U_{2} & D_{2} & D_{2} & U_{2} & U_{2} & D_{2} & D_{2} \\
U_{3} & D_{3} & U_{3} & D_{3} & U_{3} & D_{3} & U_{3} & D_{3}
\end{array}\right)
$$

It follows from the above discussion that $\mathrm{M}_{1}$ (the set of martingale measures on $\Omega_{1}$ ) is a subset of $\Delta\left(\Omega_{1}\right)$ (the set of all probability measures on $\Omega_{1}$ ) such that each $\mathbf{P} \in \mathrm{M}_{1}$ satisfies the system of linear equations

$$
\begin{equation*}
\Psi \mathbf{P}=\mathbf{R} \tag{3.3}
\end{equation*}
$$

where $\boldsymbol{\Psi}$ is given by (3.2) and $\mathbf{R}$ is a column vector of size $m$ as follows:

$$
\mathbb{R}=\left(\begin{array}{c}
R  \tag{3.4}\\
R \\
\vdots \\
R
\end{array}\right)
$$

Let $\mathbf{A} \subset \mathbb{R}^{N}$ be an affine subspace of solutions of (3.3):

$$
\mathbf{A}=\left\{\mathbf{P} \in \mathbb{R}^{N} \mid \Psi \mathbf{P}=\mathbf{R}\right\} .
$$

We can now summarize the description of the set $\mathrm{M}_{1}$ of martingale measures on $\Omega_{1}$. Geometrically $\mathrm{M}_{1}$ forms a subset of $\mathbb{R}^{N}$ which is an intersection of the simplex $\Delta\left(\Omega_{1}\right) \subset \mathbb{R}^{N}$ of all probability measures on $\Omega_{1}$ with an affine subspace $\mathbf{A}$ :

$$
\mathrm{M}_{1}=\Delta\left(\Omega_{1}\right) \cap \mathbf{A}
$$

Thus $\mathrm{M}_{1}$ is a bounded convex polytope or, in other words, the convex hull of finitely many points called vertices. Recall that the risk-neutral measures are those martingale measures $\mathbf{P}=\left(p_{1}, \ldots, p_{N}\right)$ for which $p_{j}>0$ for all $j=1, \ldots, N$. Thus the polytope $\mathrm{M}_{1}$ of martingale measures on $\Omega_{1}$ is the closure of $\mathrm{N}_{1}$, the set of risk-neutral measures on $\Omega_{1}$.

3C. Interval of no-arbitrage continent claim prices. Consider a European contingent claim $X=f(S(1))$ in the single-step model. Here $S(1)=\left(S_{1}(1), S_{2}(1), \ldots, S_{m}(1)\right)$ is the stock price vector at maturity and $f: \mathbb{R}^{m} \rightarrow \mathbb{R}$ is the pay-off function of $X$.

In this single-step model, there is only one node on the supporting tree $\mathbf{T}$. It is a root of $\mathbf{T}$ that corresponds to time $t=0$. Therefore, there is only one open interval of the no-arbitrage prices of $X$ (see (2.9)). We will use a simplified notation $\left(\mathrm{C}_{\min }(0), \mathrm{C}_{\max }(0)\right)$ for that open interval.

Recall (see Section 2B) that each no-arbitrage price of $X$ at time zero $\mathrm{C}_{\mathbf{P}}(X, 0)$ (or, equivalently, each point in the above open interval) is obtained by computing the expectation of $X$ with respect to a risk-neutral measure on $\Omega_{1}$ discounted to time zero:

$$
\begin{equation*}
\mathrm{C}_{\mathbf{P}}(X, 0)=R^{-1} \mathbf{E}_{\mathbf{P}}(X) \tag{3.5}
\end{equation*}
$$

where $\mathbf{P} \in \mathrm{N}_{1}$. Since the set $\mathrm{M}_{1}$ of martingale measures on $\Omega_{1}$ is the closure of the set $\mathrm{N}_{1}$ of risk-neutral measures on $\Omega_{1}$ we obtain that

$$
\left[\mathrm{C}_{\min }(0), \mathrm{C}_{\max }(0)\right]=\left\{\mathrm{C}_{\mathbf{P}}(X, 0) \in \mathbb{R} \mid \mathbf{P} \in \mathrm{M}_{1}\right\}
$$

It follows that the upper and lower bounds of the above interval are:

$$
\begin{align*}
\mathrm{C}_{\max }(0) & =\max \left\{\mathrm{C}_{\mathbf{P}}(X, 0) \mid \mathbf{P} \in \mathrm{M}_{1}\right\}  \tag{3.6}\\
\mathrm{C}_{\min }(0) & =\min \left\{\mathrm{C}_{\mathbf{P}}(X, 0) \mid \mathbf{P} \in \mathrm{M}_{1}\right\} . \tag{3.7}
\end{align*}
$$

Finding the quantities $\mathrm{C}_{\max }(0)$ and $\mathrm{C}_{\min }(0)$ is an important problem of contingent claim pricing.

3D. Computing bounds of the no-arbitrage contingent claim price interval via simplex algorithm. With the given contingent claim $X$ we will associate a vector $\mathbf{X}=$ $\left(X_{1}, X_{2}, \ldots, X_{N}\right) \in \mathbb{R}^{N}$, where

$$
\begin{equation*}
X_{j}=f\left(S_{1}(1)\left(\omega_{j}\right), \ldots, S_{m}(1)\left(\omega_{j}\right)\right) \tag{3.8}
\end{equation*}
$$

where $j=1, \ldots, N$. Then the expected value of $X$ with respect to a probability measure $\mathbf{P} \in \Delta\left(\Omega_{1}\right)$ can be presented as follows:

$$
\begin{equation*}
\mathbf{E}_{\mathbf{P}}(X)=\sum_{i=1}^{N} X_{i} p_{i}=\langle\mathbf{X}, \mathbf{P}\rangle \tag{3.9}
\end{equation*}
$$

where $\langle\cdot, \cdot\rangle$ denotes the Euclidean scalar product in $\mathbb{R}^{N}$.

For a fixed $X$, a map C. $(X, 0)$ is a linear functional on $\mathbb{R}^{N}$ :

$$
\begin{equation*}
\mathrm{C}_{\mathbf{P}}(X, 0)=R^{-1}\langle\mathbf{X}, \mathbf{P}\rangle \tag{3.10}
\end{equation*}
$$

So the closed interval $\left[\mathrm{C}_{\text {min }}(0), \mathrm{C}_{\max }(0)\right]$ is the image of the set $\mathrm{M}_{1}$ of martingale measures on $\Omega_{1}$ with respect to this linear functional.

As a result, the problem of finding bounds (3.6)-(3.7) for the no-arbitrage price interval of a contingent claim $X$ can be formulated in terms of the problem of finding extrema for a linear functional C. $(X, 0)$ on a convex polytope $\mathrm{M}_{1}$ in $\mathbb{R}^{N}$.

As is well known, these extrema are attained at the vertices of $\mathrm{M}_{1}$. Moreover finding both the extremal values of C. $(X, 0)$ and the vertices of $\mathrm{M}_{1}$ at which they are attained is the standard problem of Linear Programming (LP). For example, the simplex algorithm has been specifically designed to solve this type of problems.

The following describes the algorithm to compute the upper bound $\mathrm{C}_{\max }(0)$ of the interval of no-arbitrage prices for a given contingent claim $X$.

## Given data:

- Initial stock prices: $S_{1}(0), \ldots, S_{m}(0)$.
- Risk-free growth factor: $R>0$.
- Parameters of the binomial models: $\left\{D_{i}, U_{i}\right\}$, where $i=1, \ldots, m$.
- Pay-off function $f$.


## Step one: compute input data

- Compute entries $\boldsymbol{\Psi}_{i j}$ of the matrix $\boldsymbol{\Psi}$ for $i=1, \ldots, m$ and $j=1, \ldots, N=2^{m}$ (see (3.2) ).
- Compute terminal stock prices: $S_{i}(1)\left(\omega_{j}\right)=S_{i}(0) \Psi_{i j}$, where $i=1, \ldots, m$ and $j=$ $1, \ldots, N$.
- Compute the pay-off values at maturity: $X_{j}=f\left(S_{1}(1)\left(\omega_{j}\right), \ldots, S_{m}(1)\left(\omega_{j}\right)\right), j=1, \ldots, N$.


## Step two: apply simplex algorithm

Solve the following LP problem.
Maximize function:

$$
\begin{equation*}
\mathbf{P} \mapsto \mathrm{C}_{\mathbf{P}}(X, 0)=R^{-1}\langle\mathbf{X}, \mathbf{P}\rangle=R^{-1}\left(X_{1} p_{1}+\cdots+X_{N} p_{N}\right) \tag{3.11}
\end{equation*}
$$

subject to constraints:

$$
\begin{aligned}
\boldsymbol{\Psi}_{11} p_{1}+\cdots+\boldsymbol{\Psi}_{1 N} p_{N} & =R \\
& \vdots \\
\boldsymbol{\Psi}_{m 1} p_{1}+\cdots+\boldsymbol{\Psi}_{m N} p_{N} & =R \\
p_{1}+\cdots+p_{N} & =1 \\
p_{i} & \geq 0
\end{aligned}
$$

## Output:

- The upper bound $\mathrm{C}_{\max }(0)$ of the no-arbitrage price interval for $X$ at $t=0$.
- The maximal martingale measure $\mathbf{P}_{\text {max }}$, or, in other words, the vertex $\mathbf{P}_{\max }=\left(p_{1}^{*}, \ldots, p_{N}^{*}\right) \in$ $\mathrm{M}_{1}$ at which the maximum $\mathrm{C}_{\max }(0)$ is attained.

Notice that the maximum of a functional can be attained at several vertices and the algorithm will choose one of them in such a case.

The lower bound $\mathrm{C}_{\min }(0)$ of the no-arbitrage price interval for $X$ at $t=0$ can be found along similar lines with appropriate modifications in the algorithm.

Remark 3.1. In [3, Section 3], the LP problem (3.11) is solved analytically for a special case of two assets. Explicit formulas for the bounds of the no-arbitrage option price interval as well as for the corresponding extremal martingale measures are obtained.

3E. Hedging strategies and hedging portfolios. We take a position of a seller of a European contingent claim $X$ who sold $X$ at time $t=0$ for the price $\mathrm{C}_{\mathbf{P}}(X, 0)$, where $\mathbf{P}$ is some risk-neutral measure on $\Omega_{1}$, and is willing to hedge a short position in $X$ with a portfolio consisting of stock and bond.

A hedging strategy, or (equivalently) a hedging portfolio in our single-step model is a vector

$$
\left(\alpha_{1}, \ldots, \alpha_{m}, \beta\right) \in \mathbb{R}^{m+1}
$$

where $\beta$ is the number of bonds held at time 0 and $\alpha_{i}$ is a position in the $i$-th stock at time 0 . The hedging portfolio value process $V=(V(t))_{t=0,1}$ is defined by:

$$
\begin{equation*}
V(t)=\beta B(t)+\sum_{i=1}^{m} \alpha_{i} S_{i}(t) \tag{3.12}
\end{equation*}
$$

where $t=0,1$. The initial value of the hedging portfolio

$$
\begin{equation*}
V(0)=\beta+\sum_{i=1}^{m} \alpha_{i} S_{i}(0) \tag{3.13}
\end{equation*}
$$

is called a set-up cost of the hedging portfolio. If an investor is willing to invest the whole amount $\mathrm{C}_{\mathbf{P}}(X, 0)$ into a hedging strategy, then we set: $V(0)=\mathrm{C}_{\mathbf{P}}(X, 0)$.

3F. Extremal hedging strategies. Let $\left(\alpha_{1}, \ldots, \alpha_{m}, \beta\right) \in \mathbb{R}^{m+1}$ be a hedging strategy for a contingent claim $X$. We will call it a minimum cost super-hedge for $X$ if:

$$
\begin{equation*}
V(1)\left(\omega_{j}\right) \geq f\left(S_{1}(1)\left(\omega_{j}\right), \ldots, S_{m}(1)\left(\omega_{j}\right)\right)=X_{j} \tag{3.14}
\end{equation*}
$$

for all $\omega_{j} \in \Omega_{1}$, and the strategy has minimal set-up cost $V(0)$. In order to identify a minimum cost super-hedge, one needs to solve the following LP problem.

Minimize the functional

$$
\begin{equation*}
\left(\alpha_{1}, \ldots, \alpha_{m}, \beta\right) \mapsto \beta+\sum_{i=1}^{m} \alpha_{i} S_{i}(0) \tag{3.15}
\end{equation*}
$$

subject to constraints:

$$
\begin{aligned}
\beta R+\alpha_{1} S_{1}(0) \boldsymbol{\Psi}_{11}+\alpha_{2} S_{2}(0) \boldsymbol{\Psi}_{21} \cdots+\alpha_{m} S_{m}(0) \boldsymbol{\Psi}_{m 1} & \geq X_{1} \\
\beta R+\alpha_{1} S_{1}(0) \boldsymbol{\Psi}_{12}+\alpha_{2} S_{2}(0) \boldsymbol{\Psi}_{22} \cdots+\alpha_{m} S_{m}(0) \boldsymbol{\Psi}_{m 2} & \geq X_{2} \\
& \vdots \\
\beta R+\alpha_{1} S_{1}(0) \boldsymbol{\Psi}_{1 N}+\alpha_{2} S_{2}(0) \boldsymbol{\Psi}_{2 N} \cdots+\alpha_{m} S_{m}(0) \boldsymbol{\Psi}_{m N} & \geq X_{N} .
\end{aligned}
$$

It is straightforward to verify that the LP problem (3.15) is dual to the LP problem (3.11) studied in Section 3D. The next Proposition follows from the duality theorem of linear programming (see e.g. [4, Appendix]).

Proposition 3.2. Let $X=f(S(1))$ be a European contingent claim in the single-step binomial model with $m$ risky assets. There exists a minimum cost super hedging strategy $\left(\alpha_{1}, \ldots, \alpha_{m}, \beta\right) \in \mathbb{R}^{m+1}$ for $X$. This strategy satisfies (3.14) and has a set-up cost which is equal to the upper bound of the no-arbitrage contingent claim price interval:

$$
V(0)=\mathrm{C}_{\max }(0)
$$

By analogy with a minimum cost super-hedge, we will introduce a maximum cost subhedge for $X$. It is a hedging strategy $\left(\alpha_{1}, \ldots, \alpha_{m}, \beta\right) \in \mathbb{R}^{m+1}$ such that

$$
\begin{equation*}
V(1)\left(\omega_{j}\right) \leq f\left(S_{1}(1)\left(\omega_{j}\right), \ldots, S_{m}(1)\left(\omega_{j}\right)\right)=X_{j} \tag{3.16}
\end{equation*}
$$

for all $\omega_{j} \in \Omega_{1}$, and the strategy has maximal set-up cost $V(0)$.
Reasoning along the similar lines as above, one identifies a maximum cost sub-hedge as a solution of the appropriate LP problem. Applying the duality theorem of linear programming to this case, we obtain an analog of Proposition 3.2. The maximum cost sub-hedge exists and its set-up cost equals the lower bound of the no-arbitrage contingent claim price interval: $V(0)=\mathrm{C}_{\text {min }}(0)$.

In what follows, for simplicity we will call a minimum cost super-hedge an upper hedging strategy, or simply an upper hedge, we will call a maximum cost sub-hedge a lower hedging strategy, or simply a lower hedge for $X$. Notice that in general these strategies are arbitrage opportunities, as will be explained in Section 3G below.

Remark 3.3. In [3, Section 3], the LP problems for upper and lower hedges are solved analytically for a special case of two assets. Explicit formulas for the extreme hedging strategies are obtained.

3G. The upper and lower hedges are arbitrage strategies. Recall that if the number of stocks $m \geq 2$ then the market model is incomplete. This means that there are unattainable contingent claims traded on the market. For an unattainable contingent claim $X, \mathrm{C}_{\text {min }}(0)<$ $\mathrm{C}_{\max }(0)$, so that an interval of no-arbitrage prices is not empty. Moreover, for such contingent claim, there is no hedging strategy that perfectly replicates $X$, meaning that there is no strategy $\left(\alpha_{1}, \ldots, \alpha_{m}, \beta\right) \in \mathbb{R}^{m+1}$ for which $V(1)\left(\omega_{j}\right)=f\left(S_{1}(1)\left(\omega_{j}\right), \ldots, S_{m}(1)\left(\omega_{j}\right)\right)=X_{j}$ for all $\omega_{j} \in \Omega_{1}$.

Let $\left(\alpha_{1}, \ldots, \alpha_{m}, \beta\right) \in \mathbb{R}^{m+1}$ be an upper hedging strategy for an unattainable contingent claim $X=f(S(1))$. Recall that it means that $V(1)\left(\omega_{j}\right) \geq f\left(S_{1}(1)\left(\omega_{j}\right), \ldots, S_{m}(1)\left(\omega_{j}\right)\right)=X_{j}$ for all $\omega_{j} \in \Omega_{1}$, and the strategy has a set-up cost $V(0)=\mathrm{C}_{\max }(0)$. It follows that there exists at least one $\omega_{0} \in \Omega_{1}$ for which there is a strict inequality: $V(1)\left(\omega_{0}\right)>X_{0}$.

The following is an arbitrage strategy for a contingent claim seller:
(1) Sell a contingent claim at time $t=0$ for $\mathrm{C}_{\max }(0)$.
(2) Buy an upper hedging portfolio portfolio $\left(\alpha_{1}, \ldots, \alpha_{m}, \beta\right)$ at time $t=0$ for $\mathrm{C}_{\max }(0)$.
(3) The balance at maturity is $V(1)\left(\omega_{j}\right)-X_{j} \geq 0$ which is strictly positive for the scenario $\omega_{0}$.

The argument for the lower hedging strategy is analogous. This proves the following observation.

Proposition 3.4. Let $X=f(S(1))$ be an unattainable contingent claim in our single-step market. Then both the upper hedge and the lower hedge of $X$ are arbitrage strategies. The upper hedging strategy gives advantage to the contingent claim seller, while the lower hedging strategy gives advantage to the contingent claim buyer.

## 4. The multi-Step Case

In this section we consider the $n$-step market model with $m$ risky assets and a contingent claim $X$ with the pay-off function $f$.

We are going to extend the results of the previous section to this multi-step case. Let us first recall a few important facts from Section 2 and introduce some convenient notation.

4A. Quick review and necessary notation. A sample space of the $n$-step model is $\Omega_{n}$ and there is an associated supporting tree $\mathbf{T}$ in which the set of vertices at time $k \in \mathbb{T}$ is denoted by $\mathbf{T}_{k}$. We say that the set $\mathbf{T}_{n}$ consists of leaves, and each set $\mathbf{T}_{k}, k<n$ consists of nodes. Each set $\mathbf{T}_{k}, k=0,1, \ldots, n$ is in a bijective correspondence with the partition $\mathcal{P}_{k}$ of the sample space $\Omega_{n}$.

Any probability measure $\mathbf{P} \in \Delta\left(\Omega_{n}\right)$ in the $n$-step model defines a set of single-step probability measures $\mathbf{P}_{v} \in \Delta\left(\Omega_{1}\right)$ for each node $v$ of $\mathbf{T}$.

Let $\omega=\left(\omega^{1}, \ldots, \omega^{n}\right) \in \Omega_{n}$ be an element of the sample space. It defines a unique path on the tree $\mathbf{T}$ from the root to the leaf it represents. Such a path is a sequence of nodes which we denote by

$$
v_{0}(\omega), v_{1}(\omega), \ldots, v_{n}(\omega)
$$

where $v_{0}(\omega)$ is the root. Then for any measure $\mathbf{P} \in \Delta\left(\Omega_{n}\right)$ the value $\mathbf{P}(\omega)$ can be presented as follows:

$$
\begin{equation*}
\mathbf{P}(\omega)=\mathbf{P}_{v_{0}(\omega)}\left(\omega^{1}\right) \mathbf{P}_{v_{1}(\omega)}\left(\omega^{2}\right) \cdots \mathbf{P}_{v_{n-1}(\omega)}\left(\omega^{n}\right) \tag{4.1}
\end{equation*}
$$

We will use the above notation to write down a conditional expectation of a random variable $X$ on $\Omega_{n}$ with respect to a measure $\mathbf{P} \in \Delta\left(\Omega_{n}\right)$. Let $\mathcal{F}_{k}$ be a subalgebra of the algebra $2^{\Omega_{n}}$ of all subsets of $\Omega_{n}$. The conditional expectation $\mathbf{E}_{\mathbf{P}}\left(X \mid \mathcal{F}_{k}\right)$ is a random variable measurable
with respect to the algebra $\mathcal{F}_{k}$. Therefore it is determined by its values $\mathbf{E}_{\mathbf{P}}\left(X \mid \mathcal{F}_{k}\right)(v)$ on the nodes $v \in \mathbf{T}_{k}$.

Let us fix a node $v \in \mathbf{T}_{k}$ and let $\mathcal{P}(v)$ denote the set of the partition $\mathcal{P}_{k}$ which corresponds to the node $v$. It follows straightforwardly from the properties of the conditional expectation that the value $\mathbf{E}_{\mathbf{P}}\left(X \mid \mathcal{F}_{k}\right)(v)$ can be computed recursively starting from the leaves of the supporting tree $\mathbf{T}$ and, speaking informally, folding back the tree, by computing expectations with respect to suitable single-step measures:

$$
\begin{align*}
\mathbf{E}_{\mathbf{P}}\left(X \mid \mathcal{F}_{k}\right)(v) & =\sum_{\omega \in \mathcal{P}(v)} X(\omega) \frac{\mathbf{P}(\omega)}{\mathbf{P}(\mathcal{P}(v))} \\
& =\sum_{\omega \in \mathcal{P}(v)} X(\omega) \prod_{j=k+1}^{n} \mathbf{P}_{v_{j-1}(\omega)}\left(\omega^{j}\right) . \tag{4.2}
\end{align*}
$$

Here $\mathbf{P}_{v_{j}(\omega)}$ are the single-step measures that occur in (4.1).
Recall that if $\mathbf{P}$ is a risk-neutral (resp. martingale) measure on $\Omega_{n}$, then $\mathbf{P}_{v}$ is a risk-neutral (resp. martingale) measure on $\Omega_{1}$ for each node $v$ of $\mathbf{T}$, and vise versa.

4B. Extremal martingale measures. Let $X$ be a contingent claim. Given an $n$-step risk-neutral measure $\mathbf{P} \in \mathrm{N}_{n}$, a no-arbitrage price of $X$ at time $k$ with respect to $\mathbf{P}$ can be computed as the conditional expectation of $X$ with respect to an algebra $\mathcal{F}_{k}$ discounted to time $k$ :

$$
\begin{equation*}
\mathrm{C}_{\mathbf{P}}(X, k)=R^{-(n-k)} \mathbf{E}_{\mathbf{P}}\left(X \mid \mathcal{F}_{k}\right) . \tag{4.3}
\end{equation*}
$$

The time $k$ no-arbitrage price of $X$ is a random variable measurable with respect to the algebra $\mathcal{F}_{k}$. We will use the notation $\mathrm{C}_{\mathbf{P}}(X, k)(v)$ for the price of $X$ which corresponds to a node $v \in \mathbf{T}_{k}$. Using the representation (4.2) we have:

$$
\begin{equation*}
\mathrm{C}_{\mathbf{P}}(X, k)(v)=R^{-(n-k)} \sum_{\omega \in \mathcal{P}(v)} X(\omega) \prod_{j=k+1}^{n} \mathbf{P}_{v_{j-1}(\omega)}\left(\omega^{j}\right) . \tag{4.4}
\end{equation*}
$$

Varying the risk-neutral measures $\mathbf{P} \in \mathrm{N}_{n}$, we obtain at each node $v \in \mathbf{T}_{k}$ an open interval of no-arbitrage prices for a given contingent claim $X$ (see (2.9)):

$$
\left(\mathrm{C}_{\min }(v), \mathrm{C}_{\max }(v)\right)=\left\{\mathrm{C}_{\mathbf{P}}(X, k)(v) \in \mathbb{R} \mid \mathbf{P} \in \mathrm{N}_{n}\right\} .
$$

The upper and lower bounds of this interval are:

$$
\begin{align*}
\mathrm{C}_{\min }(v) & =\inf \left\{\mathrm{C}_{\mathbf{P}}(X, k)(v) \mid \mathbf{P} \in \mathrm{M}_{n}\right\}  \tag{4.5}\\
\mathrm{C}_{\max }(v) & =\sup \left\{\mathrm{C}_{\mathbf{P}}(X, k)(v) \mid \mathbf{P} \in \mathrm{M}_{n}\right\} . \tag{4.6}
\end{align*}
$$

Recall that $\mathrm{M}_{n}$ stands for the set of $n$-step martingale measures on $\Omega_{n}$ and $\mathrm{M}_{n}$ is a closure of $\mathrm{N}_{n}$.

Our goal is to identify the bounds $\mathrm{C}_{\min }(v)$ and $\mathrm{C}_{\max }(v)$ for each node $v$ on the supporting tree $\mathbf{T}$. We will start with the discussion of extremal martingale measures that produce these bounds.

Let $0 \in \mathbf{T}$ be the root and let $\mathbf{P}_{\min } \in \mathrm{M}_{n}$ and $\mathbf{P}_{\text {max }} \in \mathrm{M}_{n}$ be the extremal martingale measures for which the bounds $\mathrm{C}_{\min }(0)$ and $\mathrm{C}_{\max }(0)$ (respectively) are attained. In other words,

$$
\begin{align*}
\mathrm{C}_{\min }(0) & =R^{-n} \mathbf{E}_{\mathbf{P}_{\min }}(X)  \tag{4.7}\\
\mathrm{C}_{\max }(0) & =R^{-n} \mathbf{E}_{\mathbf{P}_{\max }}(X) . \tag{4.8}
\end{align*}
$$

The following proposition shows that the extremal measures $\mathbf{P}_{\text {min }}$ and $\mathbf{P}_{\max }$ produce the bounds of no-arbitrage price intervals for $X$ also at all other nodes of the tree.

Proposition 4.1. Let $\mathbf{P}_{\text {min }} \in \mathrm{M}_{n}$ and $\mathbf{P}_{\max } \in \mathrm{M}_{n}$ be the extremal martingale measures defined in (4.7) and (4.8) respectively for a given contingent claim $X$.

Then for any $k \in\{0,1, \ldots, n-1\}$ and any node $v \in \mathbf{T}_{k}$ the following equalities are true:

$$
\begin{align*}
\mathrm{C}_{\min }(v) & =R^{-(n-k)} \mathbf{E}_{\mathbf{P}_{\min }}\left(X \mid \mathcal{F}_{k}\right)(v)  \tag{4.9}\\
\mathrm{C}_{\max }(v) & =R^{-(n-k)} \mathbf{E}_{\mathbf{P}_{\max }}\left(X \mid \mathcal{F}_{k}\right)(v), \tag{4.10}
\end{align*}
$$

where the bounds $\mathrm{C}_{\min }(v)$ and $\mathrm{C}_{\max }(v)$ are defined in (4.5) and (4.6), respectively.

Proof. We will present the proof for $\mathbf{P}_{\max }$ leaving an analogous proof for $\mathbf{P}_{\min }$ to the reader.
Denote the measure $\mathbf{P}_{\text {max }}$ by $\mathbf{P}$ for simplicity. Suppose that there exists a different martingale measure $\mathbf{P}^{\prime} \in \mathrm{M}_{n}\left(\mathbf{P}^{\prime} \neq \mathbf{P}\right)$ and a node $v \in \mathbf{T}_{k}$ such that the upper bound $\mathrm{C}_{\max }(v)$ of the no-arbitrage price interval for $X$ at $v$ is attained for $\mathbf{P}^{\prime}$, rather than for $\mathbf{P}$. In other words, suppose that the following holds:

$$
\mathbf{E}_{\mathbf{P}}\left(X \mid \mathcal{F}_{k}\right)(v)<\mathbf{E}_{\mathbf{P}^{\prime}}\left(X \mid \mathcal{F}_{k}\right)(v)
$$

According to formula (4.2), this is equivalent to

$$
\begin{equation*}
\sum_{\omega \in \mathcal{P}(v)} X(\omega) \prod_{j=k+1}^{n} \mathbf{P}_{v_{j-1}(\omega)}\left(\omega^{j}\right)<\sum_{\omega \in \mathcal{P}(v)} X(\omega) \prod_{j=k+1}^{n} \mathbf{P}_{v_{j-1}(\omega)}^{\prime}\left(\omega^{j}\right) \tag{4.11}
\end{equation*}
$$

Let us now compute the expected value $\mathbf{E}_{\mathbf{P}}(X)$. According to our assumption (4.11), we have that

$$
\begin{aligned}
\mathbf{E}_{\mathbf{P}}(X) & =\sum_{\omega \in \Omega_{n}} X(\omega) \mathbf{P}(\omega) \\
& =\sum_{\omega \in \mathcal{P}(v)} X(\omega) \mathbf{P}(\omega)+\sum_{\omega \notin \mathcal{P}(v)} X(\omega) \mathbf{P}(\omega) \\
& =\sum_{\omega \in \mathcal{P}(v)} X(\omega) \prod_{j=1}^{n} \mathbf{P}_{v_{j-1}(\omega)}\left(\omega^{j}\right)+\sum_{\omega \notin \mathcal{P}(v)} X(\omega) \mathbf{P}(\omega) \\
& =\prod_{j=1}^{k} \mathbf{P}_{v_{j-1}(\omega)}\left(\omega^{j}\right) \sum_{\omega \in \mathcal{P}(v)} X(\omega) \prod_{j=k+1}^{n} \mathbf{P}_{v_{j-1}(\omega)}\left(\omega^{j}\right)+\sum_{\omega \notin \mathcal{P}(v)} X(\omega) \mathbf{P}(\omega) \\
& <\prod_{j=1}^{k} \mathbf{P}_{v_{j-1}(\omega)}\left(\omega^{j}\right) \sum_{\omega \in \mathcal{P}(v)} X(\omega) \prod_{j=k+1}^{n} \mathbf{P}_{v_{j-1}(\omega)}^{\prime}\left(\omega^{j}\right)+\sum_{\omega \notin \mathcal{P}(v)} X(\omega) \mathbf{P}(\omega) \\
& =\mathbf{E}_{\mathbf{P}^{\prime \prime}}(X),
\end{aligned}
$$

where $\mathbf{P}^{\prime \prime} \in \mathrm{M}_{n}$ is a martingale measure obtained by replacing each single-step measure $\mathbf{P}_{v}$ (which corresponds to the $n$-step measure $\mathbf{P}$ ) by the single-step measure $\mathbf{P}_{v}^{\prime}$ (which corresponds to the $n$-step measure $\mathbf{P}^{\prime}$ ) at the node $v$ and further at all the successors of $v$ on the tree. It follows from the above estimate that there exists a martingale measure $\mathbf{P}^{\prime \prime}$ for which

$$
R^{-n} \mathbf{E}_{\mathbf{P}^{\prime \prime}}(X)>\mathrm{C}_{\max }(0)
$$

which contradicts the assumption that $\mathrm{C}_{\max }(0)$ is the upper bound of the no-arbitrage price interval for $X$ at zero. The argument for $\mathbf{P}_{\text {min }}$ is analogous.

4 C . Computing bounds of the no-arbitrage contingent claim price interval. It is a consequence of the above proposition that the bounds $\mathrm{C}_{\min }(0)$ and $\mathrm{C}_{\max }(0)$ of the noarbitrage price interval for a given contingent claim $X$ can be computed recursively starting from the leaves of the supporting tree $\mathbf{T}$ and going backwards in time.

More specifically. For each penultimate node $v \in \mathbf{T}_{n-1}$ we solve a single-step LP problem described in Section 3D, where $v$ plays the role of a root of the corresponding single-step tree. The option payoff values are computed at the leaves adjacent to $v$ and are used as input data for the simplex algorithm described in Section 3D. Solving the single-step LP problem, we find a measure $\mathbf{P}_{\text {max }, v}$, a maximal martingale measure for this single-step problem and determine the upper bound $\mathrm{C}_{\max }(v)$ at the node $v$.

Once the upper bound of the no-arbitrage option price interval at each node of $\mathbf{T}_{n-1}$ is computed, the same procedure is applied at each node of $\mathbf{T}_{n-2}$. Consider a node $w \in \mathbf{T}_{n-2}$. This node will play the role of a root of the corresponding single-step tree. Each node $v \in \mathbf{T}_{n-1}$ adjacent to $w$ will play the role of the leaf of a single-step tree rooted at $w$. Each upper bound $\mathrm{C}_{\max }(v)$ at node $v$ will replace the corresponding option pay-off value at that node.

Continuing this way, we arrive at the root of the $n$-step tree and determine $\mathrm{C}_{\max }(0)$. The argument for the lower bound $\mathrm{C}_{\text {min }}(0)$ is analogous.

This algorithm works for a general European type contingent claim. It is, however, infeasible from a computational point of view for the multi-step models, because the number of nodes grows exponentially with the number of time steps in the model. We discuss the improvement of the algorithm in Section 5, where we show that for contingent claims with pay-off functions from a special class one can significantly reduce the computational complexity.

4D. Hedging strategies and hedging portfolios. Similar to the single-step case considered in Section 3E, we take a position of a seller of a European contingent claim $X$ who sold $X$ at time $t=0$ for the price $\mathrm{C}_{\mathbf{P}}(X, 0)$, where $\mathbf{P} \in \mathrm{M}_{n}$, and is willing to hedge a short position in $X$ with a portfolio consisting of stock and bond.

A hedging strategy or, equivalently, a hedging portfolio in our $n$-step model is a vector stochastic process

$$
\phi=\left(\alpha_{1}(t), \ldots, \alpha_{m}(t), \beta(t)\right)_{t=0,1, \ldots, n-1}
$$

where $\beta(t)$ is the number of bonds held in the portfolio over the time interval $[t, t+1)$ and $\alpha_{i}(t)$ is a position in the $i$-th stock held over the time interval $[t, t+1)$. We will assume that $\beta(t)$ as well as the vector $\alpha(t)=\left(\alpha_{1}(t), \ldots, \alpha_{m}(t)\right)$ are $\mathcal{F}_{t}-$ measurable for all $t=0,1, \ldots, n-1$.

The hedging portfolio value process $V=(V(t))_{t=0,1, \ldots, n}$ is defined by:

$$
\begin{equation*}
V(t)=\beta(t) B(t)+\sum_{i=1}^{m} \alpha_{i}(t) S_{i}(t) \tag{4.12}
\end{equation*}
$$

where $B(t)$ is a bond price at time $t$ and $S_{i}(t)$ is the $i$-th stock price at time $t$. The initial value of the hedging portfolio

$$
\begin{equation*}
V(0)=\beta(0)+\sum_{i=1}^{m} \alpha_{i}(0) S_{i}(0) \tag{4.13}
\end{equation*}
$$

is a set-up cost of the hedging portfolio. If an investor is willing to invest the whole amount $\mathrm{C}_{\mathbf{P}}(X, 0)$ into a hedging strategy, then we have: $V(0)=\mathrm{C}_{\mathbf{P}}(X, 0)$.
Let us consider the flow of capital from the time moment $t$ to the time moment $t+1$. At time $t$, the hedging portfolio $(\alpha(t), \beta(t))$ is set up. Its value is given by formula (4.12). The portfolio is held until time $t+1$ when the new stock price vector $S(t+1)=\left(S_{1}(t+1), \ldots, S_{m}(t+1)\right)$ becomes known, and a bond price changes from $B(t)$ to $B(t+1)=B(t) R$. Just prior to time $t+1$, the value of the hedging portfolio is:

$$
\beta(t) B(t+1)+\sum_{i=1}^{m} \alpha_{i}(t) S_{i}(t+1)
$$

At time $t+1$, after the new stock and bond prices are announced, the hedging portfolio is rebalanced. The number of bonds changes to $\beta(t+1)$ and the number of shares of the $i-$ th stock changes to $\alpha_{i}(t+1)$. So after rebalancing at time $t+1$, the value of the hedging portfolio becomes:

$$
\beta(t+1) B(t+1)+\sum_{\substack{i=1 \\ 17}}^{m} \alpha_{i}(t+1) S_{i}(t+1) .
$$

It is convenient to associate a hedging portfolio with a node of the supporting tree. Let $v \in \mathbf{T}_{k}$ be a node of the supporting tree at time $k$. A hedging portfolio associated with the node $v$ is a vector

$$
\phi(v)=\left(\alpha_{1}(v), \ldots, \alpha_{m}(v), \beta(v)\right) \in \mathbb{R}^{m+1}
$$

where $\beta(v)$ is the number of bonds held at node $v$ after the rebalancing at time $k$ and $\alpha_{i}(v)$ is a position in the $i$-th stock at node $v$ after the rebalancing at time $k$. The value of the hedging portfolio associated with the node $v$ after rebalancing at time $k$ is

$$
\begin{equation*}
V_{\phi(v)}(v)=\beta(v) B(k)+\sum_{i=1}^{m} \alpha_{i}(v) S_{i}(v) . \tag{4.14}
\end{equation*}
$$

Here $S_{i}(v)$ is the $i$-th stock price at node $v$ (corresponding to time $k$ ).
Let us denote by $v \omega \in \mathbf{T}_{k+1}$ a node adjacent to $v$. We will use the notation $V_{\phi(v)}(v \omega)$ for the value of the hedging portfolio associated with node $v$ just before the rebalancing at node $v \omega$ (at time $k+1$ ). We have:

$$
\begin{equation*}
V_{\phi(v)}(v \omega)=\beta(v) B(k+1)+\sum_{i=1}^{m} \alpha_{i}(v) S_{i}(v \omega) . \tag{4.15}
\end{equation*}
$$

Here $S_{i}(v \omega)$ is the $i$-th stock price at node $v \omega$ (corresponding to time $k+1$ ).
4E. Extremal hedging strategies. Assume that the maximal value $C_{\max }(v)$ of the contingent claim $X$ is computed for each node $v \in \mathbf{T}$. If follows from Proposition 3.2 that for each node $v \in \mathbf{T}$ and for each node $v \omega$ adjacent to $v$, there exists a portfolio $\phi_{\max }(v)=$ $\left(\alpha_{\max }(v), \beta_{\max }(v)\right)$ such that

$$
\begin{align*}
V_{\phi_{\max }(v)}(v) & =C_{\max }(v) \\
V_{\phi_{\max }(v)}(v \omega) & \geq C_{\max }(v \omega) . \tag{4.16}
\end{align*}
$$

The set-up cost of the hedging strategy $\phi_{\max }$ equals $C_{\max }(0)$. This strategy produces a surplus at each time step with non-zero probability. By analogy with the single-step case, this strategy is called a minimum-cost super hedge. The following proposition summarizes the above discussion.

Proposition 4.2. Let $X$ be a European type contingent claim. There exists a minimum cost super-hedging strategy $\phi_{\max }$ with the set-up cost $C_{\max }(0)$ and satisfying conditions (4.16) for each node $v \in \mathbf{T}$.

A similar argument yields a maximum-cost sub-hedge. That is, a strategy $\phi_{\min }$ with initial cost equal to $C_{\min }(0)$ and satisfying

$$
\begin{align*}
V_{\phi_{\min }(v)}(v) & =C_{\min }(v) \\
V_{\phi_{\min }(v)}(v \omega) & \leq C_{\min }(v \omega) . \tag{4.17}
\end{align*}
$$

for each node $v \in \mathbf{T}$.

## 5. Improvements of the algorithm

5A. The recombinant graph. The recursive algorithm for computing the bounds of the no-arbitrage contingent claim interval described in the previous section can be improved by descending to the recombinant graph, which we define as follows.

Let $\mathbf{T}$ be the supporting tree for our $n$-step market model. Introduce the following equivalence relation on the vertices of $\mathbf{T}$.

Definition 5.1. Let $u, v \in \mathbf{T}_{k}, 0<k \leq n$. Let $\omega_{u}$ and $\omega_{v}$ be ( $m \times k$ )-matrices associated with $u$ and $v$ respectively. Vertices $u$ and $v$ are called equivalent if and only if the sum of row entries of the matrix $\omega_{u}$ is equal to the sum of row entries of the matrix $\omega_{v}$ for each row. For a given vertex $v$, we will denote by $[v]$ an equivalence class consisting of vertices that are equivalent to $v$.

Definition 5.2. Let $\mathbf{T}_{\mathrm{r}}$ be a directed rooted graph with the vertex set consisting of the above equivalence classes and such that there is an edge from $[v]$ to $[u]$ if and only if there is an edge from $v$ to $u$ in the tree $\mathbf{T}$. The graph $\mathbf{T}_{\mathrm{r}}$ is called the recombinant graph.

Similar to the case of the supporting tree $\mathbf{T}$, we will call terminal vertices of the recombinant graph $\mathbf{T}_{r}$ leaves, and non-terminal vertices of $\mathbf{T}_{\mathrm{r}}$ nodes. Observe that each node of the recombinant graph has $2^{m}$ outgoing edges, however, the number of incoming edges can vary. Thus each node and its descendants form the tree which corresponds to the appropriate single-step model.

The recursive algorithm for computing the bounds of the no-arbitrage contingent claim intervals is essentially the same as described before and it works for a general contingent claim of a European type. For each penultimate node we solve the single-step optimization problem and proceed recursively to the root. Since the recombinant graph in the $n$-step model with $m$ assets has $(k+1)^{m}$ nodes at level $k \leq n$, running the algorithm for small $m$ and not too big $n$ becomes feasible.

Example 5.3. We have tested the above improvement by running a computer program that computes the vertices of the polytope of martingale measure, finds the maximal martingale measure and the values of $C_{\max }(v)$ for each node of the recombinant graph. For example, the program terminates within a few seconds for $m=5$ assets and $n \leq 8$. Also it takes a second to compute the above data for $m=12$ assets and $n=1$ step. This is relevant in view of the next improvement.

5B. Bounds of the no-arbitrage price interval for special contingent claims. For contingent claims that belong to a certain class (fibrewise supermodular contingent claims), further improvements of the algorithm described in Section 4C are available. For such contingent claims, both the maximal martingale measure and the minimal martingale measure are product measures. As a result, the bounds of the no-arbitrage contingent claim price interval can be computed by means of solving the LP problem described in Section 3D only once.

Theorem 5.4. Suppose a contingent claim $X: \Omega_{n} \rightarrow \mathbb{R}$ is fibrewise supermodular (see Definition A.10).
(i) There exists a single-step martingale measure $\mathbf{P} \in \mathrm{M}_{1}$ such that $\mathbf{P}_{\max }=\mathbf{P} \otimes \cdots \otimes \mathbf{P}=\mathbf{P}^{n}$. That is, the maximal martingale measure is a product measure.
(ii) If $m=2$ or $\sum_{i=1}^{m} \frac{R-D_{i}}{U_{i}-D_{i}} \leq 1$ then the minimal martingale measure is also product: $\mathbf{P}_{\text {min }}=\left(\mathbf{P}^{\prime}\right)^{n}$, for some single-step martingale measure $\mathbf{P}^{\prime}$.

Proof. (i). This part follows immediately from Theorem A.12, (i) (see Appendix), with $b_{i}=\frac{R-D_{i}}{U_{i}-D_{i}}, i=1, \ldots, m($ see (A.4)).
(ii). If $m>2$, the statement follows immediately from Theorem A.12, (ii) (see Appendix).

In the case of $m=2$ assets, the single step martingale measure form an interval and the extremal measures are its endpoints. If the maximal martingale measure in an $n$-step model is a product $\mathbf{P}_{\max }=\mathbf{P}$ then it means that $\mathbf{P}$ is an endpoint of the above interval. Thus the minima are attained on the other endpoint $\mathbf{P}^{\prime}$. It follows that $\mathbf{P}_{\min }=\left(\mathbf{P}^{\prime}\right)^{n}$.

Corollary 5.5. Under the conditions of Theorem 5.4, for a fibrewise supermodular contingent claim $X$, one has the following:

$$
\begin{align*}
\mathrm{C}_{\max }(v) & =\mathbf{E}_{\mathbf{P}^{n}}\left(X \mid \mathcal{F}_{k}\right)(v) \\
& =\sum_{\omega \in \mathcal{P}\left(\omega^{1} \ldots \omega^{k}\right)} \prod_{j=k+1}^{n} \mathbf{P}\left(\omega^{j}\right) X(\omega), \tag{5.1}
\end{align*}
$$

for $v \in \mathbf{T}_{k}$ corresponding to the set $\mathcal{P}\left(\omega^{1} \cdots \omega^{k}\right) \in \mathcal{P}_{k}$. The summation goes over the elements $\omega=\left(\omega^{1} \cdots \omega^{n}\right) \in \Omega_{n}$ with fixed first $k$ columns. Moreover,

$$
\begin{align*}
\mathrm{C}_{\min }(v) & =\mathbf{E}_{\left(\mathbf{P}^{\prime}\right)^{n}}\left(X \mid \mathcal{F}_{k}\right)(v) \\
& =\sum_{\omega \in \mathcal{P}^{\prime}\left(\omega^{1} \cdots \omega^{k}\right)} \prod_{j=k+1}^{n} \mathbf{P}^{\prime}\left(\omega^{j}\right) X(\omega), \tag{5.2}
\end{align*}
$$

In what follows we describe the maximal and the minimal martingale measure explicitly and evaluate the above formulas.
6. Pricing and hedging fibrewise supermodular contingent claims

6A. Explicit formula for the upper bound of a no-arbitrage price interval. The one-step maximal martingale measure $\mathbf{P}$ from Theorem 5.4 can be described explicitly as follows (see Equation (A.6)). Let $\mu_{k}=(1,1, \ldots, 1,0, \ldots, 0)^{\top}$ be the column vector with the first $k$ entries equal to 1 and the rest zero. Then

$$
\mathbf{P}(\omega)= \begin{cases}b_{i}-b_{i+1} & \text { if } \omega=\mu_{i}  \tag{6.1}\\ 0 & \text { otherwise }\end{cases}
$$

where $b_{i}=\frac{R-D_{i}}{U_{i}-D_{i}}$ and $b_{0}=1$. Notice that this measure is highly degenerate because among its $2^{m}$ entries only $m$ of them are nonzero.

Example 6.1. If $m=2$ then

$$
\mathbf{P}\binom{1}{1}=b_{2}, \mathbf{P}\binom{1}{0}=0, \mathbf{P}\binom{0}{1}=b_{1}-b_{2}, \mathbf{P}\binom{0}{0}=1-b_{1}
$$

which agrees with the formula for $Q_{\lambda_{+}}$in [3, Remark 1].
Thus in the formula (5.1) the summation takes place over matrices with fixed $k$ columns and the remaining columns are one of the $\mu_{i}$ 's only:

$$
\begin{equation*}
C_{\max }(v)=\sum_{i \in I^{n-k}} \mathbf{P}\left(\mu_{i_{1}}\right) \cdots \mathbf{P}\left(\mu_{i_{n-k}}\right) X\left(\omega^{1} \cdots \omega^{k} \mu_{i_{1}} \cdots \mu_{i_{n-k}}\right) \tag{6.2}
\end{equation*}
$$

where $i=\left(i_{1}, \ldots, i_{n-k}\right), i_{j} \in I=\{0,1, \ldots, m\}$ and the first $k$ columns $\omega^{1} \cdots \omega^{k}$ correspond to the vertex $v \in \mathbf{T}_{k}$. Notice that even in this case the number of summands for $C_{\max }(0)$ is exponential in $n$. However, if the payoff is path-independent then many of these summands are equal and we get:

$$
\begin{equation*}
C_{\max }(v)=\sum_{k_{0}+\ldots+k_{m}=n} \frac{n!}{k_{0}!\cdots k_{m}!} \mathbf{P}\left(\mu_{0}\right) \cdots \mathbf{P}\left(\mu_{m}\right) X(\omega^{1} \cdots \omega^{k} \underbrace{\mu_{0} \cdots \mu_{0}}_{k_{0} \text { times }} \cdots \underbrace{\mu_{m} \cdots \mu_{m}}_{k_{m} \text { times }}) \tag{6.3}
\end{equation*}
$$

6B. Explicit formula for the lower bound of a no-arbitrage price interval. Let $\nu_{i} \in \mathbb{R}^{m}$, for $i=1, \ldots, m$ be the standard basis vector. That is, the $i$-th coordinate of $\nu_{i}$ is equal to 1 and the others are zero. Suppose that $\sum_{i=1}^{m} b_{i} \leq 1$, where $b_{i}=\frac{R-D_{i}}{U_{i}-D_{i}}, i=1, \ldots, m$ that is, the assumption of Theorem 5.4, (ii) is satisfied. It follows from definition given in Equation (A.7) in the Appendix that the minimal one-step martingale measure $\mathbf{P}^{\prime}$ is then given by

$$
\mathbf{P}^{\prime}(\omega)= \begin{cases}1-\sum_{i=1}^{m} b_{i} & \text { if } \omega=(0, \ldots, 0)^{\top} \\ b_{i} & \text { if } \omega=\nu_{i} \\ 0 & \text { otherwise }\end{cases}
$$

Example 6.2. If $m=2$ we have

$$
\mathbf{P}^{\prime}\binom{1}{1}=0, \mathbf{P}^{\prime}\binom{1}{0}=b_{1}, \mathbf{P}^{\prime}\binom{0}{1}=b_{2}, \mathbf{P}^{\prime}\binom{0}{0}=1-b_{1}-b_{2},
$$

which agrees with the formula for $Q_{\lambda_{-}}$in [3, Remark 1].
The formulas for $C_{\min }(v)$ are analogous to the formulas (6.2) and (6.3) for $C_{\max }(v)$ in which the factors $\mathbf{P}\left(\mu_{i}\right)$ are replaced by $\mathbf{P}^{\prime}\left(\nu_{i}\right)$.

6C. Extremal hedging strategies. The purpose of this section is to present an effective algorithm to compute the minimum-cost superhedge and the maximum-cost subhedge for a firbewise supermodular contingent claim $X$. Let us start with the minimum-cost superhedge. It follows from Theorem 5.4 that the values of $C_{\max }(v)$ can be computed effectively for each vertex of the supporting tree; see formulas (6.2) and (6.3).
(1) Given $C_{\max }(0)$ and $C_{\max }(v)$ for each $v \in \mathbf{T}_{1}$, it follows from Proposition 3.2 that there exists a hedging portfolio $\left(\alpha_{1}(1), \ldots, \alpha_{m}(1), \beta(1)\right)$ such that its value satisfies

$$
V_{(\alpha(1), \beta(1))}(v) \geq C_{\max }(v)
$$

for each $v \in \mathbf{T}_{1}$.
(2) At time $t=1$ one of the scenarios for the first step has been realized and the maximal value of the contingent claim is $C_{\max }\left(v_{1}\right)$, where $v_{1} \in \mathbf{T}_{1}$ is the vertex corresponding to the realized scenario. If $V_{(\alpha(1), \beta(1))}\left(v_{1}\right)>C_{\max }\left(v_{1}\right)$ then the surplus (a local residual) can be withdrawn and $C_{\max }\left(v_{1}\right)$ is the value of the new hedging portfolio.
(3) Since the values $C_{\max }(w)$ for every $w \in \mathbf{T}_{2}$ adjacent to $v_{1}$ are known, we repeat the 1step argument again as in (1). That is, there exists a portfolio $\left(\alpha_{1}(2), \ldots, \alpha_{m}(2), \beta(2)\right)$ with initial value $C_{\max }\left(v_{1}\right)$ and such that

$$
V_{(\alpha(2), \beta(2))}(v) \geq C_{\max }(w)
$$

for each $w \in \mathbf{T}_{1}$ adjacent to $v_{1}$.
(4) We repeat this argument at each time until the final portfolio is chosen at time $t=n-1$.

Observe that in the above approach the dual linear program as in Proposition 3.2 is run $n$ times.

For $m=2$ it is not difficult to give explicit formulas for the hedging portfolios, see [3, Theorem 4]. For a general $m$ the task at each step is to minimize a linear functional in $\mathbf{R}^{m+1}$ on a set defined by $2^{m}$ inequalities.

The algorithm of computing a maximum-cost subhedge is similar to the above described algorithm for computing a minimum-cost superhedge and is left to the reader.

## 7. Applications of Theorem 5.4 and concrete examples

7A. Fibrewise supermodular contingent claims. In the present setting a random variable $X: \Omega_{n} \rightarrow \mathbb{R}$ is fibrewise supermodular (cf. Definition A.10) if for each $k=1,2, \ldots, n$ its restriction to the subset consisting of all entries, except those in the $k$-th column, fixed is supermodular (see definition on page 29).

In what follows we present a fairly general construction which we will subsequently specify to concrete examples of contingent claims.

Definition 7.1. A function $p: \mathbb{R}^{k} \rightarrow \mathbb{R}$ is called $\mathbb{R}$-polynomial in $k$ variables if $p\left(x_{1}, \ldots, x_{k}\right)$ is a linear combination of the Cobb-Douglas functions: $x_{1}^{p_{1}} \cdots x_{k}^{p_{k}}$, where $0 \leq p_{i} \in \mathbb{R}$; see [5, Proposition 2.2.4]. In an ordinary polynomial the exponents $p_{i}$ are non-negative integers.

Lemma 7.2. Let $h: \mathbb{R} \rightarrow \mathbb{R}$ be a convex function and let $p\left(x_{11}, \ldots, x_{m n}\right)$ be an $\mathbb{R}$-polynomial in mn variables with non-negative coefficients. Let $S_{i}(j), i=1, \ldots, m, j=1, \ldots, n$ be stock price values in the $n$-step market model with $m$ assets. Then the random variable $X: \Omega_{n} \rightarrow \mathbb{R}$

$$
X=h\left(p\left(S_{1}(1), \ldots, S_{m}(1), \ldots, S_{1}(n), \ldots, S_{m}(n)\right)\right)
$$

is fibrewise supermodular.
Proof. Since $S_{i}(k)=S_{i}(0) \psi_{i}(1) \cdots \psi_{i}(k)$, the polynomial in the stock prices $S_{i}(j)$ is also a polynomial in the stock price ratios $\psi_{i}(j)$. By restricting it to the element of the sample space $\Omega_{n}$ consisting of matrices with fixed all columns but the $k$-th one we obtain an $\mathbb{R}$-polynomial (with non-negative coefficients) in $\psi_{i}(k)$. The statement then follows from Proposition 2.2.4 (b) and Proposition 2.2.5 (a) in [5].

Example 7.3 (European basket call option). Let $K \geq 0$ and let $h(x)=(x-K)^{+}:=$ $\max \{x-K, 0\}$ and consider a random variable

$$
X=h\left(\sum_{i} a_{i} S_{i}(n)\right)=\left(\sum_{i} a_{i} S_{i}(n)-K\right)^{+}
$$

where $a_{i} \geq 0, \sum_{i} a_{i}=1$. The argument inside the function $h$ is clearly a polynomial in $S_{i}(n)$ with non-negative coefficients and hence Lemma 7.2 applies. Consequently, $X$ is fibrewise supermodular. Observe that the same conclusion follows from [5, Proposition 2.2.6]
Evaluating formula (6.3) for the upper bound of the no-arbitrage contingent claim price interval and the corresponding formula for the lower bound we obtain
$C_{\max }(v)=\sum_{k_{0}+\ldots+k_{m}=n-k} \frac{n!}{k_{0}!\cdots k_{m}!} \mathbf{P}\left(\mu_{0}\right)^{k_{0}} \ldots \mathbf{P}\left(\mu_{m}\right)^{k_{m}}\left(\sum_{i=1}^{m} a_{i} D_{i}^{d_{v}+k_{0}+\ldots+k_{i-1}} U_{i}^{u_{v}+k_{i}+\ldots+k_{m}}-K\right)^{+}$
$C_{\min }(v)=\sum_{k_{0}+\ldots+k_{m}=n-k} \frac{n!}{k_{0}!\cdots k_{m}!} \mathbf{P}^{\prime}\left(\nu_{0}\right)^{k_{0}} \cdots \mathbf{P}^{\prime}\left(\nu_{m}\right)^{k_{m}}\left(\sum_{i=1}^{m} a_{i} D_{i}^{d_{v}+k_{0}+\ldots+k_{i-1}} U_{i}^{u_{v}+k_{i}+\ldots+k_{m}}-K\right)^{+}$,
where $d_{v}, u_{v}$ are chosen so that $D_{i}^{d_{v}} U_{i}^{u_{v}}$ corresponds to the vertex $v \in \mathbf{T}_{k}$. The formula for the minimum is correct under the assumption that $\sum_{i} b_{i} \leq 1$.
By replacing $h$ with any other convex function we obtain an analogous formula for a more general contingent claim.

Example 7.4 (European basket put option). Let $h: \mathbb{R} \rightarrow \mathbb{R}$ be a convex function. Consider a random variable

$$
X=h\left(K-\sum_{i} a_{i} S_{i}(n)\right)
$$

where $K, a_{i} \geq 0$ and $\sum a_{i}=1$. If $h(x)=x^{+}$then we obtain the standard European basket put option. Notice that if $h$ is convex then so is $x \mapsto h(K-x)$. Thus by restricting the function $-\sum_{i} a_{i} S_{i}(n)$ to the elements of the sample space with all but the $k$-th column fixed we obtain a polynomial in $\psi_{i}(k)$ with non-positive coefficients. It follows from the version of Proposition 2.2.6 (a) in [5] with coefficients $a_{i} \leq 0$ (the proof is analogous) that $X$ is fibrewise supermodular and hence Theorem 5.4 applies. The formulas are similar to the ones in the previous example.

Example 7.5 (Arithmetic average Asian basket call or put option). Let

$$
X=\left(\frac{1}{n}\left(\sum_{i} a_{1 i} S_{i}(1)+\cdots+\sum_{i} a_{n i} S_{i}(n)\right)-K\right)^{+}
$$

where $a_{k i} \geq 0$ and $\sum_{i} a_{k i}=1$ for each $k$. It is thus an Asian basket call option. It follows directly form Lemma 7.2 that $X$ is fibrewise supermodular and the upper bounds of the no-arbitrage values of the contingent claim $X$ can be computed with the formula (6.2) and the analogous one for the lower bounds under the required assumption. Similarly
a put contingent claim, obtained by negating the function inside ()$^{+}$, is also fibrewise supermodular.

7B. Fibrewise submodular contingent claims. In this section we present for completeness a result for fibrewise submodular contingent claims. Since submodular contingent claims are rare we omit the proof which is a straightforward adaptation of the proof for the fibrewise supermodular case starting with Theorem A.12. Recall that a function $f$ is submodular if $-f$ is supermodular.

Theorem 7.6. Let $\mathbf{P}$ and $\mathbf{P}^{\prime}$ be the 1-step martingale probability measures from Theorem 5.4 (described precisely in Section 6A and 6B). Let $X: \Omega_{n} \rightarrow \mathbb{R}$ be a fibrewise submodular contingent claim. Then the lower bound of a no-arbitrage price interval for $X$ at a vertex $v \in \mathbf{T}_{k}$ is given by

$$
C_{\min }(v)=\mathbf{E}_{\mathbf{P}^{n}}\left(X \mid \mathcal{F}_{k}\right)(v) .
$$

If $\sum_{i} b_{i} \leq 1$ then the upper bound of the no-arbitrage price interval for $X$ at a vertex $v \in \mathbf{T}_{k}$ is equal to

$$
C_{\max }(v)=\mathbf{E}_{\left(\mathbf{P}^{\prime}\right)^{n}}\left(X \mid \mathcal{F}_{k}\right)(v)
$$

Example 7.7 (Geometric average Asian call). Let $S(k)=\sum_{i} a_{k i} S_{i}(k)$. Consider an Asian call option based on geometric mean. Its pay-off is given by

$$
X=(\sqrt[n]{S(1) \cdots S(n)}-K)^{+}
$$

This pay-off is neither super- nor submodular. However, the Arithmetic Mean - Geometric Mean Inequality (and Example 7.5) yields an upper bound of the no-arbitrage price values.

For scenarios such that $\sqrt[n]{S(1) \cdots S(n)} \geq K$ we have that $X=\sqrt[n]{S(1) \cdots S(n)}-K$ and hence, up to an additive constant, $X$ is a composition of a polynomial $S(1) \cdots S(n)$ with a concave function $\sqrt[n]{ }$. Thus Theorem 7.6 applies and the bounds of the no-arbitrage price interval can be computed effectively for some vertices $v \in \mathbf{T}$.

7C. Examples where the extremal martingale measures are not product. Consider a 1 -step model with assets $S_{1}, S_{2}: \Omega \rightarrow \mathbb{R}$ each following a binomial model with respective price ratios $0<D_{1}, D_{2}<R<U_{1}, U_{2}$, where $R$ is the risk-free rate. Let $p_{i}=\frac{R-D_{i}}{U_{i}-D_{i}}$ be the risk neutral measure for each of the asset in its own binomial model. The equations (3.1) have the form

$$
\begin{aligned}
U_{1} q_{1}+U_{1} q_{2}+D_{1} q_{3}+D_{1} q_{4} & =R \\
U_{2} q_{1}+D_{2} q_{2}+U_{2} q_{3}+D_{2} q_{4} & =R \\
q_{1}+q_{2}+q_{3}+q_{4} & =1 \\
q_{j} & \geq 0 .
\end{aligned}
$$

It is a straightforward computation that the solution set is the interval of points of the form

$$
Q(t)=\left(t, p_{1}-t, p_{2}-t, 1-p_{1}-p_{2}+t\right) \in \mathbb{R}^{4}
$$

where $\max \left\{p_{1}+p_{2}-1,0\right\}=t_{\text {min }} \leq t \leq t_{\max }=\min \left\{p_{1}, p_{2}\right\}$. In particular, for any of the choices made the interval is parallel to the vector $(1,-1,-1,1)$.

Let $X=\left(X_{1}, X_{2}, X_{3}, X_{4}\right)$ be the contingent claim. It follows that $C_{\min }(0)=\left\langle X, Q\left(t_{\min }\right)\right\rangle$ and $C_{\max }(0)=\left\langle X, Q\left(t_{\max }\right)\right\rangle$ provided that $\langle X,(1,-1,-1,1)\rangle=X_{1}-X_{2}-X_{3}+X_{4}>0$. The opposite inequality implies that that maximal and the minimal values are attained at $Q\left(t_{\min }\right)$ and $Q\left(t_{\max }\right)$, respectively. This observation is the used in the following example.

Example 7.8. Let $m=2$ and $n=2$. Consider two assets $S_{1}$ and $S_{2}$ with the following initial data:

$$
\begin{aligned}
& S_{1}(0)=100, U_{1}=1.2, D_{1}=0.8, K_{1}=100 \\
& S_{2}(0)=90, U_{1}=1.15, D_{1}=0.9, K_{2}=110
\end{aligned}
$$

Consider a spread

$$
X=\left(\frac{S_{1}(2)+S_{2}(2)}{2}-K_{1}\right)^{+}-\left(\frac{S_{1}(2)+S_{2}(2)}{2}-K_{2}\right)^{+}
$$

If the prices of both assets go up in the first step then at time $t=2$ the values of $X$ are given by

$$
X\left(\begin{array}{c}
1 \\
1 \\
1
\end{array}\right)=(10,10,7.5125,0)
$$

and when at time $t=1$ the first asset goes up and the second down we have

$$
X\left(\begin{array}{c}
1 \\
0 \\
0
\end{array}\right)=(10,8.45,0,0) .
$$

The above are straightforward calculations. The inner product with $(1,-1,-1,0)$ of the first one is negative while of the second one is positive. This shows that the single-step maximal martingale measures corresponding to these two conditional situations are distinct and hence the maximal martingale measure cannot be product of the same single-step measure.

## Appendix A.

Notions and notation. Let $\Omega$ be a finite sample space and let $\Delta(\Omega)$ be the set of all probability measures on $\left(\Omega, 2^{\Omega}\right)$. Each probability measure in $\Delta(\Omega)$ can be identified with its probability function (as was done throughout this paper), so the set $\Delta(\Omega)$ is defined as follows:

$$
\Delta(\Omega)=\left\{x: \Omega \rightarrow \mathbb{R}: \sum_{\omega \in \Omega} x(\omega)=1, x(\omega) \geq 0\right\}
$$

This is the standard simplex in the space $\mathbb{R}^{N}$, where $N$ is the number of elements in $\Omega$. Given a probability measure $x \in \Delta(\Omega)$, any function $f: \Omega \rightarrow \mathbb{R}$ is viewed as a random variable on $\left(\Omega, 2^{\Omega}, x\right)$.
We will denote by $\mathbf{E}_{x}(f)$ the expected value of $f$ with respect to the probability measure $x \in \Delta(\Omega)$.

For any $\omega \in \Omega$ we denote by $e_{\omega}: \Omega \rightarrow\{0,1\}$ the following indicator function:

$$
e_{\omega}\left(\omega^{\prime}\right)=\left\{\begin{array}{ll}
1 & \text { if } \omega^{\prime}=\omega \\
0 & \text { if } \omega^{\prime} \neq \omega
\end{array},\right.
$$

for any $\omega^{\prime} \in \Omega$. We will write $\mathbf{1}$ for the constant function on $\Omega$ : $\mathbf{1}(\omega)=1$ for any $\omega \in \Omega$.

Let $A_{1}, \ldots, A_{n}$ be finite sets and let $f_{i}: A_{i} \rightarrow \mathbb{R}$ be functions. Define the function $f_{1} \otimes \cdots \otimes f_{n}: A_{1} \times \cdots \times A_{n} \rightarrow \mathbb{R}$ by

$$
\left(f_{1} \otimes \cdots \otimes f_{n}\right)\left(a_{1}, \ldots, a_{n}\right)=f_{1}\left(a_{1}\right) \cdots f_{n}\left(a_{n}\right)
$$

where $a_{k} \in A_{k}$.
Single-step Bernoulli trials. Let us fix $m \geq 1$ and denote $\mathcal{L}=2^{\{1, \ldots, m\}}$, the power-set of $\{1, \ldots, m\}$. The set $\mathcal{L}$ is the natural sample space for $m$ single-step Bernoulli trials. Each set $S \in \mathcal{L}$ consists of numbers that correspond to trials in which a "success" occurred. For example, the set $S=\{2,5,6\}$ corresponds to the scenario where "success" occurred in trials 2,5 , and 6 , and "failure" occurred in the rest of the trials; the empty set $S=\{\emptyset\}$ corresponds to the scenario where "failure" occurred in all trials, etc.

Remark A.1. The set $\mathcal{L}$ can be put into one-to-one correspondence with the sample space $\Omega_{1}$ of a single-step market model with $m$ assets (see Section 3). Recall that $\Omega_{1}$ consists of $(m \times 1)$-matrices with binary coefficients and the number of elements in $\Omega_{1}$ is $N=2^{m}$.

Let us introduce a set of random variables $\ell_{i}: \mathcal{L} \rightarrow\{0,1\}, i=1, \ldots, m$ as follows:

$$
\ell_{i}(S)=\left\{\begin{array}{ll}
1 & \text { if } i \in S \\
0 & \text { if } i \notin S
\end{array},\right.
$$

where $S \in \mathcal{L}$. The random variable $\ell_{i}$ identifies scenarios $S \in \mathcal{L}$ in which "success" occurred in the $i$-th trial. In other words, $\ell_{i}(S)$ represents the result of the $i$-th trial in scenario $S$, where $\ell_{i}(S)=1$ corresponds to "success" and $\ell_{i}(S)=0$ corresponds to "failure". For simplicity, we will call random variables $\ell_{1}, \ldots, \ell_{m}$ the Bernoulli trials.

Remark A.2. The Bernoulli trials $\ell_{i}: \mathcal{L} \rightarrow\{0,1\}, i=1, \ldots, m$ can be put into one-to-one correspondence with the stock price ratios $\psi_{i}(1): \Omega_{1} \rightarrow\left\{D_{i}, U_{i}\right\}, i=1, \ldots, m$ (see Section 3). Recall that

$$
\psi_{i}(1)(\omega)= \begin{cases}D_{i} & \text { if } \omega_{i 1}=0 \\ U_{i} & \text { if } \omega_{i 1}=1\end{cases}
$$

It follows that

$$
\begin{equation*}
\psi_{i}(\omega)=\left(U_{i}-D_{i}\right) \ell_{i}(S)+D_{i} . \tag{A.1}
\end{equation*}
$$

(here we used the simplified notation $\psi_{i}(\omega)$ to denote $\psi_{i}(1)(\omega)$ ). In (A.1), the set $S \in \mathcal{L}$ corresponds to the sample space element $\omega \in \Omega_{1}$. For example, for $m=3, S=\{1,2\}$ corresponds to $\omega=(1,1,0)^{T}$.

It would be convenient to write

$$
\ell_{0}=1 \quad \text { and } \quad \mu_{i}=\{1, \ldots, i\} \in \mathcal{L}
$$

for every $i=0, \ldots, m$. Notice that $\mu_{0}=\emptyset$. Also observe that for all $0 \leq j \leq m$

$$
\ell_{i}\left(\mu_{j}\right)= \begin{cases}1 & \text { if } i \leq j \\ 0 & \text { if } i>j\end{cases}
$$

This identity holds for $i=0$ as well as for all $i=1, \ldots, m$. We will also denote

$$
\nu_{0}=\emptyset \quad \underset{26}{\text { and }} \quad \nu_{i}=\{i\}
$$

Clearly if $1 \leq i \leq m$ then $\ell_{i}\left(\nu_{j}\right)=\delta_{i, j}$.
Single-step martingale measures. Fix $0 \leq b_{1}, \ldots, b_{m} \leq 1$ and denote $\mathbf{b}=\left(b_{1}, \ldots, b_{m}\right)$. Define a subset $\mathrm{M}_{1}(\mathbf{b}) \subset \Delta(\mathcal{L})$ as follows:

$$
\begin{equation*}
\mathrm{M}_{1}(\mathbf{b})=\left\{x \in \Delta(\mathcal{L}): \mathbf{E}_{x}\left(\ell_{i}\right)=b_{i} \text { for all } i=1, \ldots, m\right\} . \tag{A.2}
\end{equation*}
$$

Notice that $\mathrm{M}_{1}(\mathbf{b})$ is a polyhedral set in $\mathbb{R}^{N}$, where $N=2^{m}$.
Remark A.3. If $b_{i}=\frac{R-D_{i}}{U_{i}-D_{i}}, i=1, \ldots, m$, the set $\mathrm{M}_{1}(\mathbf{b})$ coincides with the set $M_{1}$ of martingale measures on $\Omega_{1}$ (see Section 3). Indeed, consider system (3.1) which defines a martingale measure $\mathbf{P}=\left(p_{1}, p_{2}, \ldots, p_{N}\right)$ on $\Omega_{1}$ and replace each $p_{i}$ with $x_{i}$. Further, use (A.1) to express each $\psi_{i}\left(\omega_{j}\right)$ in terms of $\ell_{i}\left(S_{j}\right)$, where $S_{j}$ corresponds to $\omega_{j}$. After some straightforward algebraic transformations, (3.1) becomes:

$$
\begin{align*}
\sum_{j=1}^{N} x_{j} \ell_{i}\left(S_{j}\right) & =\frac{R-D_{i}}{U_{i}-D_{i}}  \tag{A.3}\\
\sum_{j=1}^{N} x_{j} & =1 \\
x_{i} & \geq 0
\end{align*}
$$

for $i=1, \ldots, m$. This system of equations and inequalities is equivalent to the definition (A.2) of the set $\mathrm{M}_{1}(\mathbf{b})$ for the case

$$
\begin{equation*}
b_{i}=\frac{R-D_{i}}{U_{i}-D_{i}} \tag{A.4}
\end{equation*}
$$

for $i=1, \ldots, m$.
Optimization problems 1 and 2. Given a random variable $f: \mathcal{L} \rightarrow \mathbb{R}$ we will consider the following optimization problems:

Problem 1. Find $q^{*} \in \mathrm{M}_{1}(\mathbf{b})$ such that

$$
\mathbf{E}_{q^{*}}(f)=\max \left\{\mathbf{E}_{p}(f): p \in \mathrm{M}_{1}(\mathbf{b})\right\} .
$$

Problem 2. Find $q_{*} \in \mathrm{M}_{1}(\mathbf{b})$ such that

$$
\mathbf{E}_{q_{*}}(f)=\min \left\{\mathbf{E}_{p}(f): p \in \mathrm{M}_{1}(\mathbf{b})\right\} .
$$

Remark A.4. Let (A.4) hold. Then the above optimization Problem 1 is equivalent to the LP problem (3.11) of finding the upper limit of the no-arbitrage contingent claim price interval and the maximal martingale measure in a single-step market model with $m$ assets (see Section 3D). Similarly, Problem 2 is equivalent to the LP problem of finding the lower limit of the no-arbitrage contingent claim price interval and the minimal martingale measure.

In what follows, we will always impose the following
Assumption. The vector $\mathbf{b}$ is decreasing, namely

$$
\begin{equation*}
b_{1} \geq \underset{27}{\cdots} \geq b_{m} . \tag{A.5}
\end{equation*}
$$

Recall that for a fixed $i, b_{i}=\mathbf{E}_{x}\left(\ell_{i}\right)$, where $x$ is any probability measure from the set $\mathrm{M}_{1}(\mathbf{b})$ and thus the condition (A.5) is easily fulfilled by permuting the Bernoulli trials $\ell_{i}$, $i=1, \ldots, m$.

It will be convenient to denote

$$
b_{0}=1
$$

Before we proceed with the solutions of the above optimization problems, we need to introduce some necessary definitions.

Supermodular vertex measures. Given the setup above define for $i=0, \ldots, m$

$$
q^{(i)}=b_{i}-b_{i+1}, \quad \text { and } \quad q^{(m)}=b_{m}
$$

Notice that $q^{(0)}=1-b_{1}$. Define a function $q^{*}: \mathcal{L} \rightarrow \mathbb{R}$ by

$$
\begin{equation*}
q^{*}=\sum_{i=0}^{m} q^{(i)} \cdot e_{\mu_{i}} \tag{A.6}
\end{equation*}
$$

Notice that $q^{(i)} \geq 0$ since $\mathbf{b}$ is decreasing (see (A.5)) and $0 \leq b_{i} \leq 1$. Therefore, $q^{*} \geq 0$. Also

$$
\sum_{S \in \mathcal{L}} q^{*}(S)=\sum_{i=0}^{m} q^{(i)}=\left(1-b_{1}\right)+\left(\sum_{i=1}^{m-1} b_{i}-b_{i+1}\right)+b_{m}=1
$$

So $q^{*}$ is a probability measure on $\mathcal{L}$. Also,

$$
\mathbf{E}_{q^{*}}\left(\ell_{i}\right)=\sum_{S \in \mathcal{L}} q^{*}(S) \ell_{i}(S)=\sum_{j=0}^{m} q^{(j)} \ell_{i}\left(\mu_{j}\right)=\sum_{j=i}^{m} q^{(j)}=\sum_{j=i}^{m-1}\left(b_{i}-b_{i+1}\right)+b_{m}=b_{i}
$$

for all $i=0, \ldots, m$. We deduce that

$$
q^{*} \in \mathrm{M}_{1}(\mathbf{b}) .
$$

We call $q^{*}$ the upper supermodular vertex measure of $\mathrm{M}_{1}(\mathbf{b})$. Indeed, it is a vertex of this polyhedral set, although we will not use this fact directly.

Define for all $i=0, \ldots, m$

$$
q_{(0)}=1-\sum_{i=1}^{m} b_{i}, \quad \text { and } \quad q_{(i)}=b_{i}
$$

Define a function $q_{*}: \mathcal{L} \rightarrow \mathbb{R}$ by

$$
\begin{equation*}
q_{*}=\sum_{i=0}^{m} q_{(i)} \cdot e_{\nu_{i}} . \tag{A.7}
\end{equation*}
$$

Since $b_{i} \geq 0$ it is clear that if $\sum_{i=1}^{m} b_{i} \leq 1$ then $q_{*} \geq 0$. It is also clear that

$$
\sum_{S \in \mathcal{L}} q_{*}(S)=\sum_{j=0}^{m} q_{(j)}=1
$$

Hence, $q_{*}$ is a probability measure on $\mathcal{L}$. Moreover, for any $i=1, \ldots, m$

$$
\mathbf{E}_{q_{*}}\left(\ell_{i}\right)=\sum_{S \in \mathcal{L}} \ell_{i}(S) \cdot q_{*}(S)=\sum_{j=0}^{m} \ell_{i}\left(\nu_{j}\right) \cdot q_{(j)}=b_{i}
$$

since $\ell_{i}\left(\nu_{j}\right)=\delta_{i, j}$. We deduce that

$$
q_{*} \in \mathrm{M}_{1}(\mathbf{b}) \quad \text { provided } \quad \sum_{i=1}^{m} b_{i} \leq 1
$$

We call $q_{*}$ the lower supermodular vertex measure of $\mathrm{M}_{1}(\mathbf{b})$. Indeed, it is a vertex of $\mathrm{M}_{1}(\mathbf{b})$, but we will not use this fact directly.

Supermodular functions. Recall (see e.g. [5, Section 2]) that a function $f: \mathcal{L} \rightarrow \mathbb{R}$ is called supermodular if for any $S, T \in \mathcal{L}$

$$
f(S \cup T)+f(S \cap T) \geq f(S)+f(T)
$$

We observe that the functions $\ell_{1}, \ldots, \ell_{m}$ give rise to an injective function

$$
\left(\ell_{1}, \ldots, \ell_{m}\right): \mathcal{L} \rightarrow \mathbb{R}^{m}
$$

whose image is the set $\{0,1\}^{m}$ of the vertices of the cube $[0,1]^{m}$. This gives a convenient way to define several important supermodular functions. See [5, Section 2] and [6, 7] for examples and properties of supermodular functions.

## Solving optimization Problems 1 and 2.

Theorem A.5. Suppose that $\mathbf{b}$ is decreasing (see (A.5)).
(i) For any supermodular function $f: \mathcal{L} \rightarrow \mathbb{R}$

$$
\max \left\{\mathbf{E}_{x}(f): x \in \mathrm{M}_{1}(\mathbf{b})\right\}=\mathbf{E}_{q^{*}}(f) .
$$

(ii) Suppose that $\sum_{i=1}^{m} b_{i} \leq 1$. Then for any supermodular function $f: \mathcal{L} \rightarrow \mathbb{R}$

$$
\min \left\{\mathbf{E}_{x}(f): x \in \mathrm{M}_{1}(\mathbf{b})\right\}=\mathbf{E}_{q_{*}}(f) .
$$

That is, the maximum expectation over $\mathrm{M}_{1}(\mathbf{b})$ of any supermodular function is always attained at the upper supermodular vertex measure and the minimum is attained at the lower supermodular vertex measure, if the latter is defined.

Proof. (i) Define for $i=0, \ldots, m$

$$
\alpha_{0}=f\left(\mu_{0}\right), \quad \text { and } \quad \alpha_{i}=f\left(\mu_{i}\right)-f\left(\mu_{i-1}\right)
$$

Define a function $g: \mathcal{L} \rightarrow \mathbb{R}$ by

$$
g=\sum_{i=0}^{m} \alpha_{i} \cdot \ell_{i} .
$$

Observe that for any $i=0, \ldots, m$

$$
g\left(\mu_{i}\right)=\sum_{j=0}^{m} \alpha_{j} \cdot \ell_{j}\left(\mu_{i}\right)=\sum_{j=0}^{i} \alpha_{j}=f\left(\mu_{i}\right) .
$$

Set $h=g-f$. We claim that $h \geq 0$, namely $h(S) \geq 0$ for all $S \in \mathcal{L}$. Suppose this is false. Among all $S \subseteq\{1, \ldots, m\}$ for which $h(S)<0$, choose one with the largest possible $k$ such
that $\mu_{k}=\{1, \ldots, k\} \subseteq S$. Clearly $k<m$ since we know that $h\left(\mu_{m}\right)=0$. Also, $k+1 \notin S$. Set $T=S \cup\{k+1\}$. Then $\mu_{k+1} \subseteq T$ and $S \backslash \mu_{k}=T \backslash \mu_{k+1}$. Therefore

$$
\begin{aligned}
& h(S)=g(S)-f(S)=\alpha_{0}+\sum_{i=1}^{k} \alpha_{i}+\sum_{i=k+2}^{m} \alpha_{i} \ell_{i}(S)-f(S) \\
& h(T)=g(T)-f(T)=\alpha_{0}+\sum_{i=1}^{k+1} \alpha_{i}+\sum_{i=k+2}^{m} \alpha_{i} \ell_{i}(S)-f(T)
\end{aligned}
$$

Observe that $T=S \cup \mu_{k+1}$ and that $S \cap \mu_{k+1}=\mu_{k}$. Since $f$ is supermodular,

$$
h(S)-h(T)=f(T)-f(S)-\alpha_{k+1}=f\left(S \cup \mu_{k+1}\right)-f(S)-f\left(\mu_{k+1}\right)+f\left(\mu_{k}\right) \geq 0
$$

In particular $h(T) \leq h(S)<0$ which is a contradiction to the maximality of $k$.
We will complete the proof by showing that $\mathbf{E}_{q^{*}}(f) \geq \mathbf{E}_{x}(f)$ for any $x \in \mathrm{M}_{1}(\mathbf{b})$. First, by definition of $\mathrm{M}_{1}(\mathbf{b})$ and the linearity of the expectation for any $x \in \mathrm{M}_{1}(\mathbf{b})$ we have

$$
\mathbf{E}_{x}(g)=\alpha_{0}+\sum_{i=1}^{m} \alpha_{i} b_{i}
$$

Next, we compute

$$
\begin{array}{r}
\mathbf{E}_{q^{*}}(f)=\sum_{S \subseteq\{1, \ldots, m\}} f(S) \cdot q^{*}(S)=\sum_{i=0}^{m} f\left(\mu_{i}\right) q^{(i)}=\sum_{i=0}^{m} f\left(\mu_{i}\right) \cdot\left(b_{i}-b_{i+1}\right)+f\left(\mu_{m}\right) \cdot b_{m}= \\
b_{0} f\left(\mu_{0}\right)+\sum_{i=1}^{m}\left(f\left(\mu_{i}\right)-f\left(\mu_{i-1}\right) \cdot b_{i}=\alpha_{0}+\sum_{i=1}^{m} \alpha_{i} b_{i}=\mathbf{E}_{q^{*}}(g) .\right.
\end{array}
$$

Now, since $h \geq 0$, for any $x \in \mathrm{M}_{1}(\mathbf{b})$

$$
\mathbf{E}_{x}(f)=\mathbf{E}_{x}(g-h) \leq \mathbf{E}_{x}(g)=\alpha_{0} \mathbf{E}_{x}(\mathbf{1})+\sum_{i=1}^{m} \alpha_{i} \mathbf{E}_{x}\left(\ell_{i}\right)=\alpha_{0}+\sum_{i=1}^{m} \alpha_{i} b_{i}=\mathbf{E}_{q^{*}}(g)
$$

and the proof is complete.
(ii) Define for every $i=0, \ldots, m$

$$
\alpha_{0}=f\left(\nu_{0}\right) \quad \text { and } \quad \alpha_{i}=f\left(\nu_{i}\right)-f\left(\nu_{0}\right)
$$

Define $g: \mathcal{L} \rightarrow \mathbb{R}$ by

$$
g=\alpha_{0} \ell_{0}+\sum_{i=1}^{m} \alpha_{i} \ell_{i} .
$$

Set $h=f-g$. We claim that $h \geq 0$. Assume that this is false. Choose $S \subseteq\{1, \ldots, m\}$ of the smallest possible cardinality such that $h(S)<0$. Clearly $S \neq \emptyset$ since $g\left(\nu_{0}\right)=\alpha_{0}=f\left(\nu_{0}\right)$, so $h\left(\nu_{0}\right)=0$ (and $\nu_{0}=\emptyset$ ). Choose some $k \in S$ and set $T=S \backslash\{k\}$. Then

$$
\begin{aligned}
& h(S)=f(S)-g(S)=f(S)-\alpha_{0}-\sum_{i \in S} \alpha_{i}=f(S)-\alpha_{0}-\alpha_{k}-\sum_{i \in T} \alpha_{i} \\
& h(T)=f(T)-g(T)=f(T)-\alpha_{0}-\sum_{i \in T} \alpha_{i} .
\end{aligned}
$$

Since $f$ is supermodular

$$
h(S)-h(T)=f(S)-f(T)-\alpha_{k}=f(T \cup\{k\})-f(T)-f(\{k\})-f(\emptyset) \geq 0 .
$$

It follows that $h(T) \leq h(S)<0$, contradiction to the minimality of $|S|$.
It remains to show that $\mathbf{E}_{x}(f) \geq \mathbf{E}_{q_{*}}(f)$ for any $x \in \mathrm{M}_{1}(\mathbf{b})$. By definition, for any $x \in \mathrm{M}_{1}(\mathbf{b})$ we have $\mathbf{E}_{x}(g)=\alpha_{0}+\sum_{i=1}^{m} \alpha_{i} b_{i}$. Under the hypothesis $\sum_{i=1}^{m} b_{i} \leq 1$ we have $q_{*} \in \mathrm{M}_{1}(\mathbf{b})$, so

$$
\begin{aligned}
\mathbf{E}_{q_{*}}(f)=\sum_{S \subseteq\{1, \ldots, m\}} f(S) \cdot q_{*}(S)= & \sum_{i=0}^{m} f\left(\nu_{i}\right) \cdot q_{(i)}=\left(1-\sum_{i=1}^{m} b_{i}\right) f\left(\nu_{0}\right)+\sum_{i=1}^{m} f\left(\nu_{i}\right) b_{i}= \\
& f\left(\nu_{0}\right)+\sum_{i=1}^{m}\left(f\left(\nu_{i}\right)-f\left(\nu_{0}\right)\right) b_{i}=\alpha_{0}+\sum_{i=1}^{m} \alpha_{i} b_{i}=\mathbf{E}_{q_{*}}(g) .
\end{aligned}
$$

Next, consider any $x \in \mathrm{M}_{1}(\mathbf{b})$. Since $h \geq 0$, we clearly have $\mathbf{E}_{x}(f)=\mathbf{E}_{x}(g+h) \geq \mathbf{E}_{x}(g)=$ $\mathbf{E}_{q^{*}}(g)$. This completes the proof.

Remark A.6. Part (i) of Theorem A. 5 is tightly related to Lovász extensions [1]. Indeed, $E_{q^{*}}(f)=(-f)^{L}(\mathbf{b})$ where on the right hand side is the Lovász extension of $-f$.

We identify $m$ with the set $\{1, \ldots, m\}$ and $\mathcal{L}=\{0,1\}^{m}$ with its power set. Consier a function $f:\{0,1\}^{m} \rightarrow \mathbb{R}$. The convex closure of $f$ is the function $f^{-}:[0,1] \rightarrow \mathbb{R}$ defined by

$$
f^{-}(\mathbf{x})=\min \left\{\sum_{S \subseteq m} \alpha_{S} f(S): \sum_{S \subseteq m} \alpha_{S} \cdot \mathbf{1}_{S}=\mathbf{x}, \sum_{S \subseteq m} \alpha_{S}=1, \alpha_{S} \geq 0\right\}
$$

Thus, by definition, given $\mathbf{b} \in[0,1]^{m}$ the value of $f^{-}(\mathbf{b})$ is the minimum of $E_{\alpha}(f)$ over all probability measure $\mathrm{s} \alpha$ on $\mathcal{L}$ for which $E_{\alpha}\left(\ell_{i}\right)=x_{i}$ where $\ell_{i}: \mathcal{L} \rightarrow\{0,1\}$ are projections to the $i$-th factor, namely $\alpha \in P(\mathbf{b})$.

The Lovász extension of $f$ is the function $f^{L}:[0,1]^{m} \rightarrow \mathbb{R}$ defined as follows. Given $x \in[0,1]^{m}$ write $m=\left\{k_{1}, \ldots, k_{m}\right\}$ where $x_{k_{1}} \geq x_{k_{2}} \geq \cdots \geq x_{k_{m}}$. For any $0 \leq i \leq m$ set $S_{i}=\left\{k_{1}, \ldots, k_{i}\right\} \subseteq m$. Then there are unique $\lambda_{0}, \ldots, \lambda_{m} \geq 0$ such that $\sum_{i} \lambda_{i}=1$ and $x=\sum_{i=0}^{m} \lambda_{i} \cdot \mathbf{1}_{S_{i}}$. We define

$$
f^{L}(x)=\sum_{i=0}^{m} \lambda_{i} f\left(S_{i}\right)
$$

In the notation of this paper, if $x=\mathbf{b}$ then $\mathbf{1}_{S_{i}}=\mu_{i}$ and $\lambda_{i}=q^{(i)}$. Thus, $f^{L}(b)=E_{q^{*}}(f)$. It is well known that $f^{-}=f^{L}$ if and only if $f$ is submodular [8], and notice that $f$ is submodular iff $-f$ is supermodular. Thus, $E_{q^{*}}(f)$ is the maximum expectation of $f$ with respect to probability measures $\alpha \in P(\mathbf{b})$.

Multi-step Bernoulli trials. Fix some $n \geq 1$ and consider $\mathcal{L}^{n}$. This is the natural sample space for $n$ iterations of $m$ single-step Bernoulli trials, or, equivalently, for $m n$-step Bernoulli trials.

Remark A.7. The set $\mathcal{L}^{n}$ can be put into one-to-one correspondence with the sample space $\Omega_{n}$ of the $n$-step market model with $m$ assets (see Section 2). Recall that $\Omega_{n}$ consists of ( $m \times n$ )-matrices with binary coefficients and the number of elements in $\Omega_{n}$ is $N=2^{m n}$.

For every $1 \leq k \leq n$, define the set of random variables $\ell_{i}^{k}: \mathcal{L}^{n} \rightarrow\{0,1\}, i=1, \ldots, m$ as follows:

$$
\ell_{i}^{k} \stackrel{\text { def }}{=} \underbrace{1 \otimes \cdots \otimes 1}_{k-1 \text { times }} \otimes \ell_{i} \otimes \underbrace{1 \otimes \cdots \otimes 1}_{n-k-1 \text { times }}
$$

Notice that $\ell_{i}^{k}\left(S_{1}, \ldots, S_{n}\right)$ represents the result of the $i$-th trial in the $k$-th iteration according to scenario $\left(S_{1}, \ldots, S_{n}\right): \ell_{i}^{k}\left(S_{1}, \ldots, S_{n}\right)=\ell_{i}\left(S_{k}\right)$.

Further, for every $1 \leq k \leq n$, define a vector random variable $L^{k}=\left(\ell_{1}^{k}, \ldots, \ell_{m}^{k}\right): \mathcal{L}^{n} \rightarrow$ $\{0,1\}^{m}$. Notice that $L^{k}\left(S_{1}, \ldots, S_{n}\right)$ represents the result of the $k$-th iteration of $m$ Bernoulli trials according to scenario $\left(S_{1}, \ldots, S_{n}\right)$.

Multi-step martingale measures. We fix some $\mathbf{b}=\left(b_{1}, \ldots, b_{m}\right)$ as above.
Define a subset $\mathrm{M}_{n}(\mathbf{b}) \subset \Delta\left(\mathcal{L}^{n}\right)$ as follows:

$$
\begin{equation*}
\mathrm{M}_{n}(\mathbf{b})=\left\{p \in \Delta\left(\mathcal{L}^{n}\right): \mathbf{E}_{p}\left(L^{k} \mid \mathcal{F}_{k-1}\right)=\mathbf{b}, \text { for all } k=1, \ldots, n\right\} \tag{A.8}
\end{equation*}
$$

where $\mathcal{F}_{k-1}$ is a $\sigma$-algebra generated by the random vectors $L^{1}, L^{2}, \ldots, L^{k-1}$. It may be convenient to denote the above conditional expectation by $\mathbf{E}_{p}\left(L^{k} \mid L^{1}, \ldots, L^{k-1}\right)$.
Notice that $\mathrm{M}_{n}(\mathbf{b})$ is a polyhedral set in $\mathbb{R}^{N}$ given by the following equations:

$$
\begin{equation*}
\sum_{\tau_{k}, \ldots, \tau_{n} \in \mathcal{L}} \ell_{i}\left(\tau_{k}\right) \cdot x\left(\lambda_{1}, \ldots, \lambda_{k-1}, \tau_{k}, \ldots, \tau_{n}\right)=b_{i} \cdot \sum_{\tau_{k}, \ldots, \tau_{n} \in \mathcal{L}} x\left(\lambda_{1}, \ldots, \lambda_{k-1}, \tau_{k}, \ldots, \tau_{n}\right) \tag{A.9}
\end{equation*}
$$

where $x \in \Delta\left(\mathcal{L}^{n}\right)$ and $\left.1 \leq k \leq n, \lambda_{1}, \ldots, \lambda_{k-1} \in \mathcal{L}, i=1, \ldots, m\right)$. To see this divide both sides of (A.9) by the sum in the right hand side and observe that the resulting left hand side is exactly the required conditional expectation. Notice that the product probability measure $q \otimes \cdots \otimes q$ belongs to $\mathrm{M}_{n}(\mathbf{b})$ for $q \in \mathrm{M}_{1}(\mathbf{b})$.

Remark A.8. If (A.4) holds for $i=1, \ldots, m$, the set $\mathrm{M}_{n}(\mathbf{b})$ coincides with the set $\mathrm{M}_{n}$ of martingale measures on $\Omega_{n}$ (see (2.4) and (2.5) ).

Optimization problems 3 and 4. Given a random variable $f: \mathcal{L}^{n} \rightarrow \mathbb{R}$, we will consider the following optimization problems:

Problem 3. Find $p^{*} \in \mathrm{M}_{n}(\mathbf{b})$ such that

$$
\mathbf{E}_{p^{*}}(f)=\max \left\{\mathbf{E}_{p}(f): p \in \mathrm{M}_{n}(\mathbf{b})\right\} .
$$

Problem 4. Find $p_{*} \in \mathrm{M}_{n}(\mathbf{b})$ such that

$$
\mathbf{E}_{p_{*}}(f)=\min \left\{\mathbf{E}_{p}(f): p \in \mathrm{M}_{n}(\mathbf{b})\right\} .
$$

Remark A.9. Let (A.4) hold. Then the above optimization Problem 3 is equivalent to the problem of finding the upper limit of the no-arbitrage contingent claim price interval at time zero and the maximal martingale measure in the $n$-step market model with $m$ assets (see Section 4C). Similarly, Problem 4 is equivalent to the problem of finding the lower limit of the no-arbitrage contingent claim price interval at time zero and the minimal martingale measure.

## Fibrewise supermodular functions.

Definition A.10. We say that $f: \mathcal{L}^{n} \rightarrow \mathbb{R}$ is fibrewise supermodular if its restriction to the subset $\left(\lambda_{1}, \ldots, \lambda_{k-1}\right) \times \mathcal{L} \times\left(\lambda_{k+1}, \ldots, \lambda_{n}\right)$ is a supermodular function on $\mathcal{L}$ for every $\lambda_{1}, \ldots, \widehat{\lambda_{k}}, \ldots, \lambda_{n} \in \mathcal{L}$.

Example A.11. Suppose that $u_{i, j}: \mathcal{L} \rightarrow \mathbb{R}$ are affine functions, $1 \leq i \leq n$ and $1 \leq j \leq r$. Suppose that each $u_{i, j}$ has the form $\sum_{k} \alpha_{k} x_{k}+\gamma$ where $\alpha_{k} \geq 0$. Suppose that $h: \mathbb{R} \rightarrow \mathbb{R}$ is convex. Let $g: \mathcal{L}^{n} \rightarrow \mathbb{R}$ be the function $\sum_{j=1}^{r} u_{1, j} \otimes \cdots \otimes u_{n, j}+c \mathbf{1}$. Then $f=\left.h \circ g\right|_{\mathcal{L}^{n}}$ is fibrewise supermodular.

Indeed, the restriction of $g$ to any subset $\left(\lambda_{1}, \ldots, \lambda_{k-1}\right) \times \mathcal{L} \times\left(\lambda_{k+1}, \ldots, \lambda_{n}\right)$ is a linear combination of the affine maps $u_{k, 1}, \ldots, u_{k, r}$ with non-negative coefficients and a multiple of $\mathbf{1}$. Now appeal to [5, Proposition 2.2.6 (a)].

## Solving optimization Problems 3 and 4.

Theorem A.12. Suppose that $\mathbf{b}$ is decreasing (see (A.5)).
(i) For any fibrewise supermodular function $f: \mathcal{L}^{n} \rightarrow \mathbb{R}$

$$
\max \left\{\mathbf{E}_{p}(f): p \in \mathrm{M}_{n}(\mathbf{b})\right\}=\mathbf{E}_{p^{*}}(f)=\mathbf{E}_{q^{*} \otimes n}(f),
$$

where $q^{*}$ is the upper supermodular vertex measure defined in (A.6).
(ii) Suppose $\sum_{i=1}^{m} b_{i} \leq 1$. Then for any fibrewise supermodular function $f: \mathcal{L}^{n} \rightarrow \mathbb{R}$

$$
\min \left\{\mathbf{E}_{p}(f): p \in \mathrm{M}_{n}(\mathbf{b})\right\}=\mathbf{E}_{p_{*}}(f)=\mathbf{E}_{q_{*} \otimes n}(f)
$$

where $q_{*}$ is the lower supermodular vertex measure defined in (A.7).
Thus, $\mathbf{E}_{p}(f)$ is maximized at the product measure $p^{*}=q^{* \otimes n} \in \mathrm{M}_{n}(\mathbf{b})$, where $q^{*}$ is the upper supermodular vertex measure (see (A.6)).

Further, $\mathbf{E}_{p}(f)$ is minimized at the product measure $p_{*}=q_{*}{ }^{\otimes n} \in \mathrm{M}_{n}(\mathbf{b})$, where $q_{*}$ is the lower supermodular vertex measure (see (A.7)), provided the latter is defined.

Proof. We will show that if $x \in \mathrm{M}_{n}(\mathbf{b})$ then $\mathbf{E}_{x}(f) \leq \mathbf{E}_{q^{*} \otimes n}(f)$ and $\mathbf{E}_{x}(f) \geq \mathbf{E}_{q_{*} \otimes n}(f)$, thus proving the result since $q^{\otimes n} \in \mathrm{M}_{n}(\mathbf{b})$ for any $q \in \mathrm{M}_{1}(\mathbf{b})$.

For every $0 \leq k \leq n$ and every $\lambda_{1}, \ldots, \lambda_{k} \in \mathcal{L}$ set

$$
y_{k}\left(\lambda_{1} \ldots \lambda_{k}\right)=\sum_{\theta_{k+1}, \ldots, \theta_{n} \in \mathcal{L}} x\left(\lambda_{1}, \ldots, \lambda_{k}, \theta_{k+1}, \ldots, \theta_{n}\right) .
$$

Thus, $y_{k}\left(\lambda_{1} \ldots \lambda_{k}\right)$ is the probabilty (with respect to $x$ ) of the event $\left\{L^{1}=\lambda_{1}, \ldots, L^{k}=\lambda_{k}\right\}$. Define $x^{(k)}: \mathcal{L}^{n} \rightarrow \mathbb{R}$ and $x_{(k)}: \mathcal{L}^{n} \rightarrow \mathbb{R}$ by

$$
\begin{aligned}
x^{(k)}\left(\lambda_{1}, \ldots, \lambda_{n}\right) & =y_{k}\left(\lambda_{1} \ldots \lambda_{k}\right) \cdot q^{*}\left(\lambda_{k+1}\right) \cdots q^{*}\left(\lambda_{n}\right) \\
x_{(k)}\left(\lambda_{1}, \ldots, \lambda_{n}\right) & =y_{k}\left(\lambda_{1} \ldots \lambda_{k}\right) \cdot q_{*}\left(\lambda_{k+1}\right) \cdots q_{*}\left(\lambda_{n}\right) .
\end{aligned}
$$

Clearly, $x^{(k)} \geq 0$ and $x_{(k)} \geq 0$. Also, since $q^{*}$ is a probability measure on $\mathcal{L}$

$$
\begin{aligned}
\sum_{\lambda_{1}, \ldots, \lambda_{n} \in \mathcal{L}} x^{(k)}\left(\lambda_{1}, \ldots, \lambda_{n}\right) & =\sum_{\lambda_{1}, \ldots, \lambda_{k}} y_{k}\left(\lambda_{1}, \ldots, \lambda_{k}\right) \sum_{\lambda_{k+1}, \ldots, \lambda_{n}} \prod_{i=k+1}^{n} q^{*}\left(\lambda_{i}\right) \\
& =\sum_{\lambda_{1}, \ldots, \lambda_{k} \in \mathcal{L}} y_{k}\left(\lambda_{1}, \ldots, \lambda_{k}\right)=\sum_{\lambda_{1}, \ldots, \lambda_{n} \in \mathcal{L}} x\left(\lambda_{1}, \ldots, \lambda_{n}\right)=1 .
\end{aligned}
$$

So $x^{(k)} \in \Delta\left(\mathcal{L}^{n}\right)$. An identical argument using $q_{*} \in \Delta(\mathcal{L})$ gives $x_{(k)} \in \Delta(\mathcal{L})$.
Clearly $x^{(n)}=x=x_{(n)}$ and $x^{(0)}=q^{* \otimes n}$ and $x_{(0)}=q_{*}{ }^{\otimes n}$. To complete the proof it remains to prove that $\mathbf{E}_{x^{(k)}}(f) \geq \mathbf{E}_{x^{(k+1)}}(f)$ for all $0 \leq k<n$, and similarly that $\mathbf{E}_{x_{(k)}}(f) \leq \mathbf{E}_{x_{(k+1)}}(f)$. If $y_{k}\left(\lambda_{1}, \ldots, \lambda_{k}\right)>0$ then the $\operatorname{map} p: \mathcal{L} \rightarrow \mathbb{R}$

$$
p(\lambda)=\frac{y_{k+1}\left(\lambda_{1}, \ldots, \lambda_{k}, \lambda\right)}{y_{k}\left(\lambda_{1}, \ldots, \lambda_{k}\right)}
$$

is clearly a probability measure on $\mathcal{L}$. Direct computation shows that for any $\lambda_{1}, \ldots, \lambda_{k} \in \mathcal{L}$

$$
\begin{aligned}
& \sum_{\tau_{k+1}, \ldots, \tau_{n} \in \mathcal{L}} \ell\left(\tau_{k+1}\right) x^{(k+1)}(\boldsymbol{\lambda}, \boldsymbol{\tau})=\sum_{\tau} y_{k+1}(\boldsymbol{\lambda} \tau) \ell_{i}(\tau) \quad \text { and } \\
& \sum_{\tau_{k+1}, \ldots, \tau_{n} \in \mathcal{L}} x^{(k+1)}(\boldsymbol{\lambda}, \boldsymbol{\tau})=\sum_{\tau} y_{k}(\boldsymbol{\lambda})
\end{aligned}
$$

The defining equations (A.9) of $\mathrm{M}_{n}(\mathbf{b})$ imply that

$$
\sum_{\tau} y_{k+1}\left(\lambda_{1}, \ldots, \lambda_{k}, \tau\right) \ell_{i}(\tau)=b_{i} \cdot y_{k}\left(\lambda_{1}, \ldots, \lambda_{k}\right)
$$

Therefore, if $y_{k}\left(\lambda_{1}, \ldots, \lambda_{k}\right)>0$ then $p \in \mathrm{M}_{1}(\mathbf{b})$. Hence, if $g: \mathcal{L} \rightarrow \mathbb{R}$ is supermodular (resp. submodular) then $\mathbf{E}_{q^{*}}(g) \geq \mathbf{E}_{p}(g)$ (resp. $\left.\mathbf{E}_{q^{*}}(g) \leq \mathbf{E}_{p}(g)\right)$ which can be written explicitly

$$
\begin{align*}
& \sum_{\tau \in \mathcal{L}} g(\tau) \cdot y_{k+1}\left(\lambda_{1}, \ldots, \lambda_{k}, \tau\right) \leq y_{k}\left(\lambda_{1}, \ldots, \lambda_{k}\right) \cdot \sum_{\tau \in \mathcal{L}} g(\tau) \cdot q^{*}(\tau)  \tag{A.10}\\
& \sum_{\tau \in \mathcal{L}} g(\tau) \cdot y_{k+1}\left(\lambda_{1}, \ldots, \lambda_{k}, \tau\right) \geq y_{k}\left(\lambda_{1}, \ldots, \lambda_{k}\right) \cdot \sum_{\tau \in \mathcal{L}} g(\tau) \cdot q_{*}(\tau) .
\end{align*}
$$

For $\theta_{1}, \ldots, \theta_{j} \in \mathcal{L}$ denote $q^{*}(\boldsymbol{\theta})=q^{*}\left(\theta_{1}\right) \cdots q^{*}\left(\theta_{j}\right)$ and similarly for $q_{*}(\boldsymbol{\theta})$. Now,

$$
\begin{aligned}
& \mathbf{E}_{x^{(k)}}(f)=\sum_{\lambda_{1}, \ldots, \lambda_{k}} \sum_{\theta_{k+1}, \ldots, \theta_{n}} f(\boldsymbol{\lambda} \boldsymbol{\theta}) \cdot y_{k}(\boldsymbol{\lambda}) \cdot q^{*}(\boldsymbol{\theta}) \quad \text { and } \\
& \mathbf{E}_{x^{(k+1)}}(f)=\sum_{\lambda_{1}, \ldots, \lambda_{k+1}} \sum_{\theta_{k+2}, \ldots, \theta_{n}} f\left(\boldsymbol{\lambda \boldsymbol { \theta } ) \cdot y _ { k + 1 } ( \boldsymbol { \lambda } ) \cdot q ^ { * } ( \boldsymbol { \theta } )}\right.
\end{aligned}
$$

Similar formulas hold for $\mathbf{E}_{x_{(k)}}(f)$ and $\mathbf{E}_{x_{(k+1)}}(f)$ by replacing $q^{*}$ with $q_{*}$ on the right hand sides. Since $f$ is fibrewise supermodular and $q^{*}(\boldsymbol{\theta}) \geq 0$ we can use the inequalities in (A.10)
to continue the second equality:

$$
\begin{aligned}
&=\sum_{\lambda_{1}, \ldots, \lambda_{k}} \sum_{\theta_{k+2}, \ldots, \theta_{n}} \sum_{\tau \in \mathcal{L}} f(\boldsymbol{\lambda} \tau \boldsymbol{\theta}) \cdot y_{k+1}(\boldsymbol{\lambda} \tau) \cdot q^{*}(\boldsymbol{\theta}) \\
& \leq \sum_{\lambda_{1}, \ldots, \lambda_{k}} \sum_{\theta_{k+2}, \ldots, \theta_{n}} \sum_{\tau \in \mathcal{L}} f(\boldsymbol{\lambda} \tau \boldsymbol{\theta}) \cdot y_{k}(\boldsymbol{\lambda}) \cdot q^{*}(\tau) \cdot q(\boldsymbol{\theta}) \\
&=\sum_{\lambda_{1}, \ldots, \lambda_{k}} \sum_{\theta_{k+1}, \ldots, \theta_{n}} f(\boldsymbol{\lambda} \boldsymbol{\theta}) \cdot y_{k}(\boldsymbol{\lambda}) \cdot q^{*}(\boldsymbol{\theta})=\mathbf{E}_{x^{(k)}}(f) .
\end{aligned}
$$

A similar calculation shows that $\mathbf{E}_{x_{(k+1)}}(f) \geq \mathbf{E}_{x_{(k)}}(f)$. This completes the proof.

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